

Given a sample x , find its label y .

$$y' = \arg \max_y P(y|x) \leftarrow \text{Discriminative model.}$$

$$= \arg \max_y \frac{P(x|y) \cdot P(y)}{P(x)}$$

$$= \arg \max_y P(x|y) \cdot P(y). \leftarrow \text{Generative model.}$$

- Discriminative Model
 - Regression : Linear Regression : $y|x \sim N(h(\phi), \sigma^2)$
 - $P(y|x; \theta)$ } Binary Classification : Logistic Regression : $y|x \sim \text{Bern}(h_{\phi}(x))$
 - Multi-classification : Softmax Regression : $y|x \sim \text{Multi}(h_{\phi}(x))$.

▷ LR : Given x , we have: $\hat{y} = (\theta^*)^T x + b^*$

▷ Classification : Given x , $h_{\phi}(x) = \begin{bmatrix} P(y=1|x) \\ \vdots \\ P(y=k|x) \end{bmatrix}$.

e.g. $h_{\phi}(x) = \begin{bmatrix} 0.1 \\ 0.2 \\ 0.7 \end{bmatrix} \quad \begin{cases} y=0 \\ y=1 \\ y=2 \end{cases} \quad , \quad y' = \arg \max_y P(Y=y|x).$

- Generative Model :
 - Continuous Input : GDA
 - $P(x,y) = P(x|y) \cdot P(y)$.
 - Discrete Input : NB
 - $P(x_1, \dots, x_n|y) = \prod_{i=1}^n P(x_i|y)$.
 - $y \sim \text{Bern}(\phi)$
 - $x_i|y=b \sim \text{Bern}(\phi_i|y=b)$.

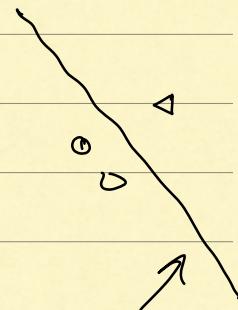
- \uparrow Statistical View Points.

- \downarrow Optimization.

We want a standard / criteria
to tell me $y = -1, 1$.

} Draw a line

} After transform, to draw a line.



Details:

① Linear Regression:

- $h(x; \theta, b) = \theta^T x + b$, $\theta, x \in \mathbb{R}^n$, $b, y \in \mathbb{R}$.

$$\begin{aligned} J(\theta) &= \frac{1}{2} \sum_{i=1}^m (h(x^{(i)}) - y^{(i)})^2 \\ &= \frac{1}{2} (x\theta - y)^T (x\theta - y). \end{aligned}$$

- Normal Equation: $\nabla_{\theta} J(\theta) = x^T x \theta - \underset{\downarrow}{x^T y}$

$$\theta^* = \underline{(x^T x)^{-1}} \cdot x^T y.$$

$$(x^T x + \alpha I)^{-1} \cdot x^T y. : \text{Ridge Reg}$$

- Gradient Descent: $\theta' := \theta - \alpha \cdot \nabla J(\theta)$.

$$\theta_1 = \theta_0 - \alpha \cdot \nabla J(\theta_0).$$

$$\theta_2 = \theta_1 - \alpha \cdot \nabla J(\theta_1)$$

$$= \underbrace{\theta_0}_{\theta} - \alpha \cdot \nabla J(\theta_0) - \alpha \cdot \nabla J(\underbrace{\theta_1}_{\theta}).$$

- Newton's Method.

- Relationship with MLE

• MLE : maximize $\prod_{i=1}^m P(y^{(i)} | x^{(i)})$, log-likelihood.

• Assumption : $y^{(i)} = \theta^T x^{(i)} + \varepsilon^{(i)}$

where $\varepsilon^{(i)} \stackrel{i.i.d.}{\sim} N(0, \sigma^2)$.
 \downarrow
 $y|x$.

• LR = MLE + Linear Model + Gaussian Assumption.

- ② Logistic Regression : $y = 0, 1$.

→ hypothesis function.

• $y|x \sim \text{Bern}(h_\theta(x))$.

$$\text{i.e. } P(y=1|x) = h_\theta(x) = (h_\theta(x))^1 (1-h_\theta(x))^{1-1}$$

$$\text{where } h_\theta(x) \triangleq g(\theta^T x) \triangleq \frac{1}{1+e^{-\theta^T x}} \quad \begin{matrix} \uparrow \\ \text{sigmoid function} \end{matrix} \quad \star$$

- ③ Softmax Regression : $y = 0, 1, \dots, k-1$.

• $y|x \sim \text{Multinomial}(h_\theta(x))$.

$$\text{where } h_\theta(x) = \begin{bmatrix} P(y=0|x) \\ \vdots \\ P(y=k-1|x) \end{bmatrix} = \frac{1}{\sum_{j=0}^{k-1} e^{\theta_j^T x}} \begin{bmatrix} e^{\theta_0^T x} \\ \vdots \\ e^{\theta_{k-1}^T x} \end{bmatrix}$$

$$v \geq 1 \quad v(z) > 1 \quad \begin{matrix} \uparrow \\ \text{normalize term}, \quad 1 > h_\theta(x) > 0 \end{matrix}$$

$$\begin{bmatrix} v_{(1)} \\ \vdots \\ v_{(k)} \end{bmatrix} \geq 1$$

$$\begin{aligned} \ell(\Theta) &= \sum_{i=1}^m \log P(y^{(i)} | x^{(i)}). \\ &= \sum_{i=1}^m \log \prod_{l=0}^{k-1} P(y^{(i)} = l | x^{(i)})^{1\{y^{(i)}=l\}}. \end{aligned}$$

$$= \sum_{i=1}^m \sum_{l=0}^{k-1} 1\{y^{(i)} = l\} \cdot \log \frac{e^{\theta_l^T x^{(i)}}}{\sum_{j=0}^{k-1} e^{\theta_j^T x^{(i)}}}$$

$$\begin{aligned}
 \nabla_{\theta_1}(\ell(\Theta)) &= \nabla_{\theta_1} \sum_{i=1}^m \sum_{e=0}^{k-1} 1\{y^{(i)} = e\} \cdot \log \frac{e^{\theta_e^\top x^{(i)}}}{\sum_{j=0}^{k-1} e^{\theta_j^\top x^{(i)}}} \\
 &= \sum_{i=1}^m \nabla_{\theta_1} \sum_{e=0}^{k-1} 1\{y^{(i)} = e\} \left(\theta_e^\top x^{(i)} - \log \sum_{j=0}^{k-1} e^{\theta_j^\top x^{(i)}} \right) \\
 &= \sum_{i=1}^m \nabla_{\theta_1} \left(\sum_{e=0}^{k-1} 1\{y^{(i)} = e\} \cdot \theta_e^\top x^{(i)} - \sum_{e=0}^{k-1} 1\{y^{(i)} = e\} \log \sum_{j=0}^{k-1} e^{\theta_j^\top x^{(i)}} \right) \\
 &= \sum_{i=1}^m \nabla_{\theta_1} \left(\sum_{e=0}^{k-1} 1\{y^{(i)} = e\} \cdot \theta_e^\top x^{(i)} - \log \sum_{j=0}^{k-1} e^{\theta_j^\top x^{(i)}} \right)
 \end{aligned}$$

$$= \nabla_{\theta_i} \sum_{l=0}^{k-1} 1\{y^{(i)} = l\} \cdot \theta_i^\top \cdot x^{(i)}$$

$$= \nabla_{\theta_1} \left(\underbrace{1\{y^{(1)} = 0\}}_{\text{Indicator function}} \cdot \theta_0^T \cdot x^{(1)} + \dots + 1\{y^{(r)} = k\}}_{\text{Indicator function}} \cdot \theta_k^T \cdot x^{(r)} \right)$$

$$= \nabla_{\theta_i} \cdot \mathbb{1}\{y^{(i)} = 1\} \cdot \theta_i^\top \cdot x^{(i)}$$

$$= 1 \cdot \{y^{(i)} = 1\} \cdot x^{(i)}$$

o-

- Relation with Logistic Regression.

$$k=2.$$

$$P_{Y|X}(1|X) = \frac{e^{\theta_1^\top X}}{e^{\theta_0^\top X} + e^{\theta_1^\top X}} = \frac{1}{e^{(\theta_0 - \theta_1)^\top X} + 1} = \frac{1}{\sigma((\theta_1 - \theta_0)^\top X)}.$$

④ Gaussian Discriminative Analysis (GDA) / QDA.

$$\cdot P(x, y) = P(x|y) \cdot P(y).$$

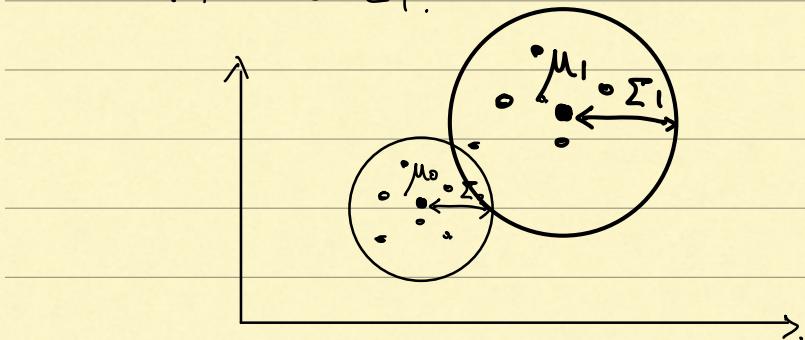
Assume $y \sim \text{Bern}(\phi)$, $\phi \in \mathbb{R}$

$$x|y=0 \sim \mathcal{N}(\mu_0, \Sigma_0), \mu_0, \Sigma_0 \in \mathbb{R}^n$$

$$x|y=1 \sim \mathcal{N}(\mu_1, \Sigma_1), \Sigma_0, \Sigma_1 \in \mathbb{R}^{n \times n}$$

$$\cdot \text{GDA: } \Sigma_0 = \Sigma_1,$$

$$\text{QDA: } \Sigma_0 \neq \Sigma_1.$$



Intuitions:

$$\textcircled{1} \quad Y \quad \begin{array}{c|c|c|c} Y & 0 & 1 \\ \hline P(Y) & \frac{1}{10} & \frac{9}{10} \end{array}$$

ϕ : how many samples
are drawn

② are drawn Datasets.

μ_1 : \mathcal{Q}_1 : the mean of samples which are labeled 1.

$$\Sigma_1: (\mathbf{x}^{(i)} - \mu_1)(\mathbf{x}^{(i)} - \mu_1)^T$$

$$\textcircled{1} \quad \phi = \frac{\sum_{i=1}^m 1 \{ y^{(i)} = 1 \}}{m}$$

$$\mu_1 = \frac{\sum_{i=1}^m 1 \{ y^{(i)} = 1 \} \cdot \mathbf{x}^{(i)}}{\sum_{i=1}^m 1 \{ y^{(i)} = 1 \}}$$

$$\Sigma_1 = \frac{\sum_{i=1}^m 1 \{ y^{(i)} = 1 \} \cdot (\mathbf{x}^{(i)} - \mu_1)(\mathbf{x}^{(i)} - \mu_1)^T}{\sum_{i=1}^m 1 \{ y^{(i)} = 1 \}}$$

$$\begin{aligned}
 l(\phi, \mu_0, \mu_1, \Sigma_0, \Sigma_1) &\stackrel{\text{def}}{=} \log \prod_{i=1}^m P(x^{(i)}, y^{(i)}) \\
 &= \sum_{i=1}^m \log P(x^{(i)}, y^{(i)}) \\
 &= \sum_{i=1}^m (\log P(x^{(i)}|y^{(i)}) + \log P(y^{(i)}))
 \end{aligned}$$

① Do one thing in one line.

$$y \sim \text{Bern}(\phi) \rightarrow P(y) = \phi^y (1-\phi)^{1-y}, \quad y=0,1.$$

$$\log P(y) = y \log \phi + (1-y) \log (1-\phi).$$

$$x|y=0 \sim N(\mu_0, \Sigma_0) \rightarrow P(x|y=0) = [(2\pi)^{\frac{n}{2}} |\Sigma_0|^{\frac{1}{2}}]^{-1} \cdot \exp(-\frac{1}{2}(x-\mu_0)^T \Sigma_0^{-1} (x-\mu_0))$$

$$\begin{aligned}
 \log P(x|y=0) &= -\frac{n}{2} \log(2\pi) - \frac{1}{2} \log |\Sigma_0| \\
 &\quad - \frac{1}{2} (x-\mu_0)^T \Sigma_0^{-1} (x-\mu_0).
 \end{aligned}$$

$$\begin{aligned}
 x|y=b \sim N(\mu_b, \Sigma_b) \Rightarrow \log P(x|y=b) &= -\frac{n}{2} \log(2\pi) - \frac{1}{2} \log |\Sigma_b| \\
 &\quad - \frac{1}{2} (x-\mu_b)^T \Sigma_b^{-1} (x-\mu_b).
 \end{aligned}$$

where $b=0,1$.

② If you want to derive the equation for all cases,

try to figure out a special case.

e.g. WA 1.4. $\nabla_{\Theta_b} l(\Theta) \rightarrow \nabla_{\Theta_1} l(\Theta)$.

$$\begin{aligned}
 l &= \sum_{i=1}^m (\log P(x^{(i)}|y^{(i)}) + \log P(y^{(i)})) \\
 &= \sum_{i=1}^m -\frac{n}{2} \log(2\pi) - \frac{1}{2} \log |\Sigma_{y^{(i)}}| - \frac{1}{2} (x^{(i)} - \mu_{y^{(i)}})^T \Sigma_{y^{(i)}}^{-1} (x^{(i)} - \mu_{y^{(i)}}) \\
 &\quad + \sum_{i=1}^m y^{(i)} \log \phi + (1-y^{(i)}) \log (1-\phi) \\
 &= \sum_{i=1}^m \sum_{b=0}^1 \mathbb{1}\{y^{(i)}=b\} \left[-\frac{n}{2} \log(2\pi) - \frac{1}{2} \log |\Sigma_b| - \frac{1}{2} (x^{(i)} - \mu_b)^T \Sigma_b^{-1} (x^{(i)} - \mu_b) \right] \\
 &\quad + \sum_{i=1}^m y^{(i)} \log \phi + (1-y^{(i)}) \log (1-\phi)
 \end{aligned}$$

$$\bullet \frac{\partial l}{\partial \phi} = 0, \quad \frac{\partial l}{\partial \mu_b} = 0, \quad \frac{\partial l}{\partial \Sigma_b} = 0, \quad b=0,1$$

$$\begin{aligned}\bullet \frac{\partial l}{\partial \phi} &= \frac{\partial}{\partial \phi} \sum_{i=1}^m 1\{y^{(i)} = 0\} \log \phi + (1-y^{(i)}) \log (1-\phi) \\ &= \sum_{i=1}^m y^{(i)} \cdot \frac{1}{\phi} + (1-y^{(i)}) \frac{-1}{1-\phi} = 0\end{aligned}$$

$$\Rightarrow \phi^* = \frac{1}{m} \sum_{i=1}^m 1\{y^{(i)} = 0\}.$$

$$\bullet \frac{\partial l}{\partial \mu_0} = \frac{\partial}{\partial \mu_0} \sum_{i=1}^m \sum_{b=0}^1 1\{y^{(i)} = b\} \left[-\frac{1}{2} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_b)^T \cdot \boldsymbol{\Sigma}_b^{-1} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_b) \right]$$

$$= \frac{\partial}{\partial \mu_0} \sum_{i=1}^m 1\{y^{(i)} = 0\} \left[-\frac{1}{2} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0)^T \cdot \boldsymbol{\Sigma}_0^{-1} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0) \right]$$

$$+ \frac{\partial}{\partial \mu_0} \sum_{i=1}^m 1\{y^{(i)} = 1\} \left[-\frac{1}{2} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_1)^T \cdot \boldsymbol{\Sigma}_1^{-1} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_1) \right]$$

$$= \frac{\partial}{\partial \mu_0} \sum_{i=1}^m 1\{y^{(i)} = 0\} \left[-\frac{1}{2} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0)^T \cdot \boldsymbol{\Sigma}_0^{-1} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0) \right]$$

$$= \frac{\partial}{\partial \mu_0} \sum_{i=1}^m 1\{y^{(i)} = 0\} \cdot \frac{\partial}{\partial \mu_0} \left[-\frac{1}{2} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0)^T \cdot \boldsymbol{\Sigma}_0^{-1} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0) \right] \quad (\star)$$

Recall: $\frac{\partial}{\partial v} (v-w)^T A (v-w) = (A+A^T)(v-w)$

$$(\star) = \sum_{i=1}^m 1\{y^{(i)} = 0\} \cdot \left[-\left(\boldsymbol{\Sigma}_0^{-1} (\boldsymbol{\mu}_0 - \mathbf{x}^{(i)}) \right) \right].$$

$$= - \boldsymbol{\Sigma}_0^{-1} \cdot \left(\sum_{i=1}^m 1\{y^{(i)} = 0\} \cdot (\boldsymbol{\mu}_0 - \mathbf{x}^{(i)}) \right) = 0.$$

$$\rightarrow \sum_{i=1}^m 1\{y^{(i)} = 0\} \cdot (\boldsymbol{\mu}_0 - \mathbf{x}^{(i)}) = 0$$

$$\rightarrow \boldsymbol{\mu}_0^* = \frac{\sum_{i=1}^m 1\{y^{(i)} = 0\} \cdot \mathbf{x}^{(i)}}{\sum_{i=1}^m 1\{y^{(i)} = 0\}}$$

$$\begin{aligned} \frac{\partial \ell}{\partial \Sigma_0} &= \frac{\partial}{\partial \Sigma_0} \cdot \sum_{i=1}^m \sum_{b=0}^1 \mathbb{1}\{y^{(i)}=b\} \left[-\frac{1}{2} \log |\Sigma_b| - \frac{1}{2} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_b)^T \cdot \Sigma_b^{-1} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_b) \right] \\ &= \sum_{i=1}^m \mathbb{1}\{y^{(i)}=0\} \cdot \frac{\partial}{\partial \Sigma_0} \left[-\frac{1}{2} \log |\Sigma_0| - \frac{1}{2} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0)^T \cdot \Sigma_0^{-1} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0) \right] \end{aligned}$$

Note that $\frac{\partial}{\partial \Sigma} \log |\Sigma| = \frac{1}{|\Sigma|} \cdot \frac{\partial |\Sigma|}{\partial \Sigma} = \frac{1}{|\Sigma|} \cdot |\Sigma| \cdot \Sigma^{-1} = \Sigma^{-1}$

$$\cdot \frac{\partial}{\partial \Sigma} \mathbf{V}^T \Sigma^{-1} \Sigma = - \Sigma^{-1} \cdot \mathbf{V} \mathbf{V}^T (\Sigma^{-1})^T.$$

$$\begin{aligned} (\star\star) \Leftrightarrow &= \sum_{i=1}^m \mathbb{1}\{y^{(i)}=0\} \cdot \left[-\frac{1}{2} \cdot \Sigma_0^{-1} + \frac{1}{2} \Sigma_0^{-1} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0) (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0)^T \cdot \Sigma_0^{-1} \right] \\ &\quad = 0. \\ \Sigma_0^* &= \frac{\sum_{i=1}^m \mathbb{1}\{y^{(i)}=0\} \cdot (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0) (\mathbf{x}^{(i)} - \boldsymbol{\mu}_0)^T}{\sum_{i=1}^m \mathbb{1}\{y^{(i)}=0\}}. \end{aligned}$$

- ① Treat notation clearly
- ② One thing at One Step.
- ③ Too many notations? Try a special case.
- ④ Burden will \downarrow , when the diff are known.
- ⑤ Turn your intuition radar on?
e.g. scalar/ vector? Make sense?

- Relation to Logistic Regression.

$$\cdot P(y=1|x) = \frac{1}{1+e^{-\Theta^T x}}$$



Lec 4. P23/28.

- If $x/y \sim N(\mu, \Sigma)$, $P(y/x)$ is a logistic function.

⑤ Naive Bayes.

- Dictionary

$$x = \begin{bmatrix} 0 \\ \vdots \\ 1 \\ 1 \\ \vdots \\ 0 \end{bmatrix} \begin{array}{l} \text{Review} \\ \text{Session} \end{array}$$

} large.

spam? not spam?

$$\begin{aligned} p(x_1, \dots, x_n | y) &= p(x_1 | y) \cdots p(x_n | y) \\ &= \prod_{i=1}^n p(x_i | y) \end{aligned}$$

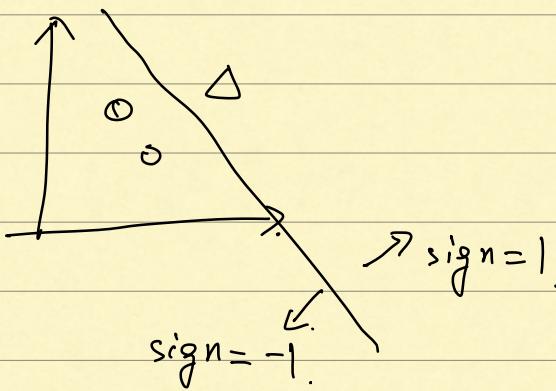
- $y \sim \text{Bern}(\phi)$

$$x_i | y=b \sim \text{Bern}(\phi_i | y=b), b=1, 2.$$

- $\phi^* = \frac{\sum_{i=1}^m 1 \{ y^{(i)} = 1 \}}{m}$

$$\phi_{j|y=1}^* = \frac{\sum_{i=1}^m 1 \{ y^{(i)} = 1 \} \cdot 1 \{ x_j^{(i)} = 1 \}}{\sum_{i=1}^m 1 \{ y^{(i)} = 1 \}}$$

⑥ SVM.



- Given $\{(x^{(i)}, y^{(i)})\}_{i=1}^m$, find (w^*, b^*) to.

prime problem : $\min_{w,b} \frac{1}{2} \|w\|^2$

$$\text{s.t. } y^{(i)}(w^T x^{(i)} + b) \geq 1, i=1, \dots, m.$$

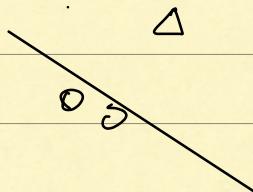
dual problem : $\max_{\alpha} \sum_{i=1}^m \alpha_i - \frac{1}{2} \sum_{i,j=1}^m \alpha_i \alpha_j y^{(i)} y^{(j)} \langle x^{(i)}, x^{(j)} \rangle$

$$\text{s.t. } \alpha_i \geq 0, i=1, \dots, m$$

$$\sum \alpha_i y^{(i)} = 0.$$

$\Rightarrow \alpha^*$

$$\text{Solution : } \begin{cases} w^* = \sum_i \alpha_i^* y^{(i)} x^{(i)} \\ b^* = -\frac{1}{2} \left(\max_{i:y^{(i)}=1} w^T x^{(i)} + \min_{i:y^{(i)}=-1} (w^*)^T x^{(i)} \right) \end{cases}$$



Given new sample z

$$y = \text{sign}[w^T z + b^*] = \text{sign}\left[\sum_{i=1}^m \alpha_i y^{(i)} \langle x^{(i)}, z \rangle + b^*\right]$$

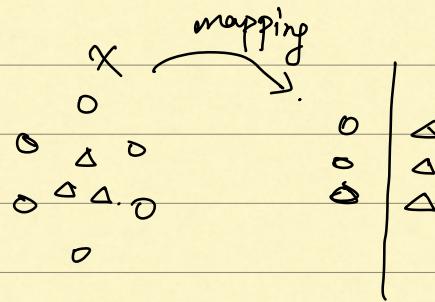
$$\text{where } \text{sign}(t) = \begin{cases} 1 & , t > 0 \\ 0 & , t = 0 \\ -1 & , t < 0. \end{cases}$$

- Soft-SVM Review at home \star .

• Kernel Trick.

Messy Data.

$$\phi: \mathbb{R}^d \rightarrow \mathbb{R}^D.$$



$$k(x, x') \triangleq \phi(x)^\top \phi(x') \in \mathbb{R}. \quad k \in \mathbb{R}^{m \times m}$$

dual problem: $\max_{\alpha} \sum_{i=1}^m \alpha_i - \frac{1}{2} \sum_{i,j=1}^m \alpha_i \alpha_j y^{(i)} y^{(j)} k(x^{(i)}, x^{(j)})$
 s.t. $\sum \alpha_i \geq 0, i = 1, \dots, m$
 $\sum \alpha_i y^{(i)} = 0.$

$$\Rightarrow \alpha^*$$

Solution : $\begin{cases} w^* = \sum_i \alpha_i^* y^{(i)} \phi(x^{(i)}) \\ b^* = y^{(j)} - \sum_{i=1}^m \alpha_i^* y^{(i)} k(x^{(i)}, x^{(j)}) \text{ for some } j \end{cases}$

• $f(x) = w^T \phi(x) + b^*$
 $= \sum_{i=1}^m \alpha_i y^{(i)} \langle \phi(x^{(i)}), \phi(x) \rangle + b.$

• $y' = \arg \max_y P(y|x).$

$y = \text{sign}(_)$

$= \arg \max_y P(x|y) \cdot P(y).$