

Learning from Data

Lecture 9: Principal Component Analysis

Yang Li yangli@sz.tsinghua.edu.cn

TBSI

November 26, 2021

Today's Lecture

Unsupervised Learning (Part II): PCA

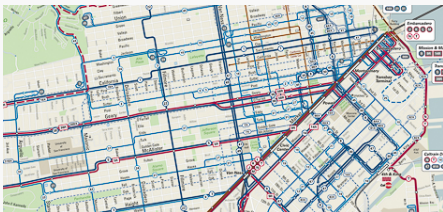
- ▶ Motivation
- ▶ Linear PCA
- ▶ Kernel PCA

Project Information: <http://yangli-feasibility.com/home/classes/lf2021fall/project.html>

Motivation

Motivation of PCA

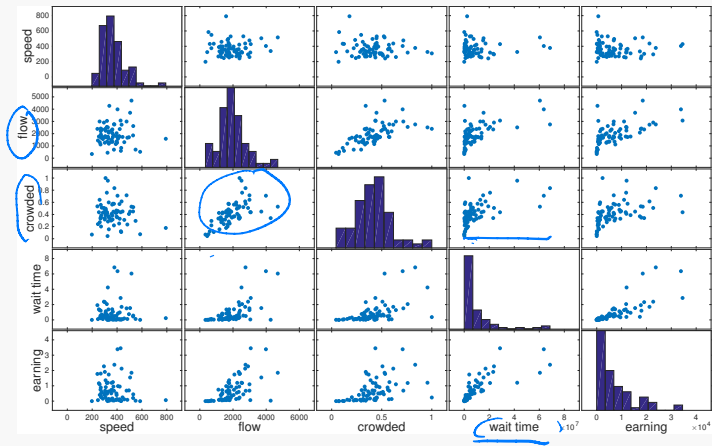
Example: Analyzing San Francisco public transit route efficiency



features	notes
speed	average speed
<u>flow</u>	# boarding passengers per hour
<u>crowded</u>	% passenger capacity reached
wait time	average waiting time at bus stop
earning	net operation revenue
⋮	⋮

Motivation of PCA

Input features contain a lot of redundancy

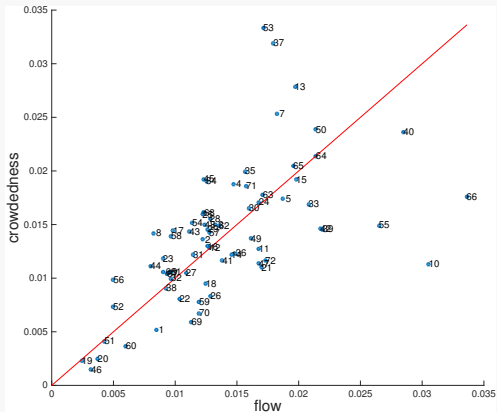


Scatter plot matrix reveals pairwise correlations among 5 major features

Motivation of PCA

Example of linearly dependent features

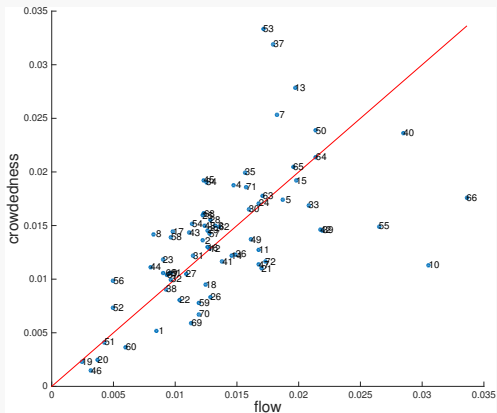
- ▶ Flow: average # boarding passengers per hour
- ▶ Crowdedness: $\frac{\text{average \# passengers on train}}{\text{train capacity}}$



Motivation of PCA

Example of linearly dependent features

- ▶ Flow: average # boarding passengers per hour
- ▶ Crowdedness: $\frac{\text{average \# passengers on train}}{\text{train capacity}}$



How can we automatically detect and remove this redundancy?

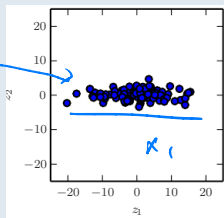
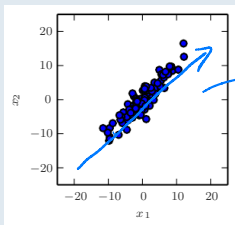
- ▶ geometric approach *start here!*
- ▶ diagonalize covariance matrix approach

How to removing feature redundancy?

Given $\{x^{(1)}, \dots, x^{(m)}\}$, $x^{(i)} \in \mathbb{R}^n$.

- ▶ Find a linear, orthogonal transformation $W: \mathbb{R}^n \rightarrow \mathbb{R}^k$ of the input data
- ▶ W aligns the direction of maximum variance with the axes of the new space.

Example: $n = 2$



features x_1 and x_2 are strongly correlated

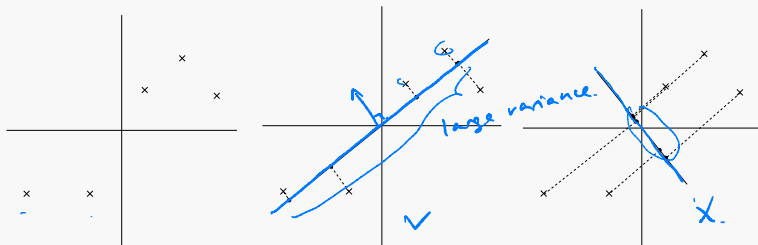
variations in $z = x^T W$ is mostly along the x -axis. x can be represented in 1D!

Direction of Maximum Variance

- Suppose $\mu = \text{mean}(x) = 0$, $\sigma_j = \text{var}(x_j) = 1$ (variance of j th feature)

Direction of Maximum Variance

- Suppose $\mu = \text{mean}(x) = 0$, $\sigma_j = \text{var}(x_j) = 1$ (variance of j th feature)
- Find **major axis of variation** unit vector u :



input observations

projections on u
have large variance

projections on u
have small variance

u maximizes the variance of the projections

Linear PCA

Principal Component Analysis (PCA)

Pearson, K. (1901), Hotelling, H. (1933) "Analysis of a complex of statistical variables into principal components". Journal of Educational Psychology.

PCA goals

- ▶ Find principal components $\underline{u}_1, \dots, \underline{u}_n$ that are mutually orthogonal (uncorrelated)
- ▶ Most of the variation in x will be accounted for by k principal components where $k \ll n$.
 $\leftarrow x \in \mathbb{R}^n$

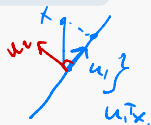
Principal Component Analysis (PCA)

Pearson, K. (1901), Hotelling, H. (1933) "Analysis of a complex of statistical variables into principal components". Journal of Educational Psychology.

PCA goals

- ▶ Find principal components u_1, \dots, u_n that are mutually orthogonal (uncorrelated)
- ▶ Most of the variation in x will be accounted for by k principal components where $k \ll n$.

Main steps of (full) PCA:



1. Standardize x such that $Mean(x) = 0$, $Var(x_j) = 1$ for all j
2. Find projection of x , $u_1^T x$ with maximum variance
3. For $j = \underline{2}, \dots, \underline{n}$,
Find another projection of x , $u_j^T x$ with maximum variance,
where u_j is orthogonal to u_1, \dots, u_{j-1}

Step 1: Standardize data

Normalize x such that $Mean(x) = 0$ and $Var(x_j) = 1$ ↗ σ_j

$$x^{(i)} := x^{(i)} - \underline{\mu} \leftarrow \text{recenter}$$

$$\rightarrow \boxed{x_j^{(i)}} := x_j^{(i)} / \underline{\sigma_j} \leftarrow \text{scale by } \text{stdev}(x_j)$$

Check:

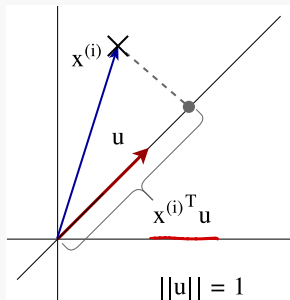
$$\text{var} \left(\frac{x_j}{\sigma_j} \right) = \frac{1}{m} \sum_{i=1}^m \left(\frac{x_j^{(i)} - \underline{\mu_j}}{\sigma_j} \right)^2 = \frac{1}{\sigma_j^2} \frac{1}{m} \sum_{i=1}^m (x_j^{(i)} - \mu_j)^2$$

$$= \frac{1}{\sigma_j^2} \sigma_j^2 = 1$$

σ_j^2

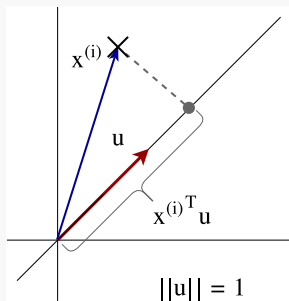
Step 2: Find Projection with Maximum Variance

Since $\|u\| = 1$, the length of $x^{(i)}$'s projection on u is $x^{(i)T}u$.



Step 2: Find Projection with Maximum Variance

Since $\|u\| = 1$, the length of $x^{(i)}$'s projection on u is $x^{(i)T}u$.



$$\frac{1}{m} \sum_{i=1}^m x^{(i)} = \mathbf{0}.$$

length of the.

Variance of the projections: $x^{(1)}, \dots, x^{(m)}$

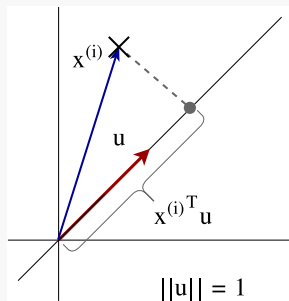
$$\frac{1}{m} \sum_{i=1}^m (x^{(i)T}u - \mathbf{0})^2 = \frac{1}{m} \sum_{i=1}^m u^T x^{(i)} x^{(i)T} u$$

Step 2: Find Projection with Maximum Variance

$$X = \begin{bmatrix} -x^{(1)} \\ -x^{(2)} \\ \vdots \\ -x^{(m)} \end{bmatrix} \quad \text{cov}(X) = \begin{bmatrix} b_{11}^2 & b_{12}^2 & \dots & b_{1n}^2 \\ & b_{22}^2 & \dots & \vdots \\ & & \dots & \vdots \\ & & & b_{nn}^2 \end{bmatrix}$$

(n x n)

Since $\|u\| = 1$, the length of $x^{(i)}$'s projection on u is $x^{(i)T}u$.



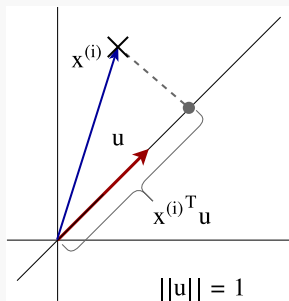
Variance of the projections:

$$\begin{aligned} \frac{1}{m} \sum_{i=1}^m (x^{(i)T}u - \mathbf{0})^2 &= \frac{1}{m} \sum_{i=1}^m \underline{u^T} x^{(i)} x^{(i)T} \underline{u} \\ &= u^T \left(\underbrace{\frac{1}{m} \sum_{i=1}^m x^{(i)} x^{(i)T}}_{\Sigma = \text{cov}(X)} \right) u \end{aligned}$$

Step 2: Find Projection with Maximum Variance

$$\Sigma = \text{cov}(X) = \frac{1}{n} X X^T.$$

Since $\|u\| = 1$, the length of $x^{(i)}$'s projection on u is $x^{(i)T} u$.



Variance of the projections:

$$\begin{aligned} \frac{1}{m} \sum_{i=1}^m (x^{(i)T} u - \mathbf{0})^2 &= \frac{1}{m} \sum_{i=1}^m u^T x^{(i)} x^{(i)T} u \\ &= u^T \left(\frac{1}{m} \sum_{i=1}^m x^{(i)} x^{(i)T} \right) u \\ &= u^T \Sigma u \end{aligned}$$

Σ : the sample covariance matrix of $x^{(1)} \dots x^{(m)}$.

1st Principal Component

Find unit vector u_1 that maximizes variance of projections:

$$u_1 = \underset{u: \|u\|=1}{\operatorname{argmax}} u^T \Sigma u \quad (1)$$

Handwritten notes: A red arrow points from the text "cov(X)" above to the Σ in the equation. A red underline is drawn under the constraint $u: \|u\|=1$.

u_1 is the **1st principal component** of X

u_1 can be solved using optimization tools, but it has a more efficient solution:

Proposition 1

u_1 is the **largest eigenvector** of covariance matrix Σ

Proposition 1

u_1 is the largest eigenvector of covariance matrix Σ

Proof.



Proposition 1

u_1 is the largest eigenvector of covariance matrix Σ

Proof. Generalized Lagrange function of Problem [1](#):

$$L(u) = -u^T \Sigma u + \beta(u^T u - 1)$$

Proposition 1

u_1 is the largest eigenvector of covariance matrix Σ

Proof. Generalized Lagrange function of Problem **1**:

$$L(u) = -u^T \Sigma u + \beta(u^T u - 1)$$

To minimize $L(u)$,

$$\frac{\delta L}{\delta u} = -2\Sigma u + 2\beta u = 0 \implies \Sigma u = \beta u$$

Therefore u_1 must be an eigenvector of Σ .

Proposition 1

u_1 is the largest eigenvector of covariance matrix Σ

Proof. Generalized Lagrange function of Problem **1**:

$$L(u) = -u^T \Sigma u + \beta(u^T u - 1)$$

To minimize $L(u)$,

$$\frac{\delta L}{\delta u} = -2\Sigma u + 2\beta u = 0 \implies \Sigma u = \beta u$$

Therefore u_1 must be an eigenvector of Σ .

Let $u_1 = v_j$, the eigenvector with the j th largest eigenvalue λ_j ,

$$u_1^T \Sigma u_1 = v_j^T \Sigma v_j = \lambda_j v_j^T v_j = \lambda_j.$$

Hence $u_1 = v_1$, the eigenvector with the largest eigenvalue λ_1 . □

Proposition 2

The j th principal component of X , u_j is the j th largest eigenvector of Σ .

Proof.

Proposition 2

The j th principal component of X , u_j is the j th largest eigenvector of Σ .

Proof. Consider the case $j = 2$,

$$u_2 = \operatorname{argmax}_{u: \|u\|=1, u_1^T u=0} u^T \Sigma u \quad (2)$$

Proposition 2

The j th principal component of X , u_j is the j th largest eigenvector of Σ .

Proof. Consider the case $j = 2$,

$$u_2 = \operatorname{argmax}_{u: \|u\|=1, u_1^T u=0} u^T \Sigma u \quad (2)$$

The Lagrangian function:

$$L(u) = -u^T \Sigma u + \beta_1 (u^T u - 1) + \beta_2 (u_1^T u)$$

Minimizing $L(u)$ yields:

$$\beta_2 = 0, \Sigma u = \beta_1 u$$

Proposition 2

The j th principal component of X , u_j is the j th largest eigenvector of Σ .

Proof. Consider the case $j = 2$,

$$u_2 = \operatorname{argmax}_{u: \|u\|=1, u_1^T u=0} u^T \Sigma u \quad (2)$$

The Lagrangian function:

$$L(u) = -u^T \Sigma u + \beta_1 (u^T u - 1) + \beta_2 (u_1^T u)$$

Minimizing $L(u)$ yields:

$$\beta_2 = 0, \Sigma u = \beta_1 u$$

To maximize $u^T \Sigma u = \lambda$, u_2 must be the eigenvector with the second largest eigenvalue $\beta_1 = \lambda_2$. The same argument can be generalized to cases $j > 2$. (Use induction to prove for $j = 1 \dots n$) □

Summary

We can solve PCA by solving an eigenvalue problem!

Main steps of (full) PCA:

1. Standardize x such that $Mean(x) = 0$, $Var(x_j) = 1$ for all j
2. Compute $\Sigma = cov(x)$
3. Find principal components u_1, \dots, u_n by eigenvalue decomposition:
 $\Sigma = U\Lambda U^T$. $\leftarrow U$ is an orthogonal basis in \mathbb{R}^n

Next we project data vectors x to this new basis, which spans the **principal component space**.

PCA Projection

- ▶ Projection of sample $x \in \mathbb{R}^n$ in the principal component space:

$$z^{(i)} = \begin{bmatrix} x^{(i)T} u_1 \\ \vdots \\ x^{(i)T} u_n \end{bmatrix} \in \mathbb{R}^n$$

PCA Projection

- ▶ Projection of sample $x \in \mathbb{R}^n$ in the principal component space:

$$z^{(i)} = \begin{bmatrix} x^{(i)T} u_1 \\ \vdots \\ x^{(i)T} u_n \end{bmatrix} \in \mathbb{R}^n$$

- ▶ Matrix notation:

$$z^{(i)} = \begin{bmatrix} | & & | \\ u_1 & \dots & u_n \\ | & & | \end{bmatrix}^T x^{(i)} = U^T x^{(i)}, \text{ or } Z = XU$$

PCA Projection

- ▶ Projection of sample $x \in \mathbb{R}^n$ in the principal component space:

$$z^{(i)} = \begin{bmatrix} x^{(i)T} u_1 \\ \vdots \\ x^{(i)T} u_n \end{bmatrix} \in \mathbb{R}^n$$

- ▶ Matrix notation:

$$z^{(i)} = \begin{bmatrix} | & & | \\ u_1 & \dots & u_n \\ | & & | \end{bmatrix}^T x^{(i)} = U^T x^{(i)}, \text{ or } Z = XU$$

- ▶ The truncated transformation $Z_k = XU_k$ keeping only the first k principal components is used for **dimension reduction**.

Properties of PCA

- ▶ The variance of principal component projections are

$$\text{Var}(x^T u_j) = u_j^T \Sigma u_j = \lambda_j \text{ for } j = 1, \dots, n$$

Properties of PCA

- ▶ The variance of principal component projections are

$$\text{Var}(x^T u_j) = u_j^T \Sigma u_j = \lambda_j \text{ for } j = 1, \dots, n$$

- ▶ % of variance explained by the j th principal component: $\frac{\lambda_j}{\sum_{i=1}^n \lambda_i}$.
i.e. projections are uncorrelated

Properties of PCA

- ▶ The variance of principal component projections are

$$\text{Var}(x^T u_j) = u_j^T \Sigma u_j = \lambda_j \text{ for } j = 1, \dots, n$$

- ▶ % of variance explained by the j th principal component: $\frac{\lambda_j}{\sum_{i=1}^n \lambda_i}$.
i.e. projections are uncorrelated
- ▶ % of variance accounted for by retaining the first k principal components ($k \leq n$): $\frac{\sum_{j=1}^k \lambda_j}{\sum_{j=1}^n \lambda_j}$

Another geometric interpretation of PCA is minimizing projection residuals. (see homework!)

Covariance Interpretation of PCA

PCA removes the “redundancy” (or noise) in input data X :

Let $Z = XU$ be the PCA projected data,

$$\text{cov}(Z) = \frac{1}{m} Z^T Z = \frac{1}{m} (XU)^T (XU) = U^T \left(\frac{1}{m} X^T X \right) U = U^T \Sigma U$$

Covariance Interpretation of PCA

PCA removes the “redundancy” (or noise) in input data X :

Let $Z = XU$ be the PCA projected data,

$$\text{cov}(Z) = \frac{1}{m} Z^T Z = \frac{1}{m} (XU)^T (XU) = U^T \left(\frac{1}{m} X^T X \right) U = U^T \Sigma U$$

Since Σ is symmetric, it has real eigenvalues. Its eigen decomposition is

$$\Sigma = U \Lambda U^T$$

where

$$U = \begin{bmatrix} | & & | \\ u_1 & \dots & u_n \\ | & & | \end{bmatrix}, \Lambda = \begin{bmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_n \end{bmatrix}$$

Covariance Interpretation of PCA

PCA removes the “redundancy” (or noise) in input data X :

Let $Z = XU$ be the PCA projected data,

$$\text{cov}(Z) = \frac{1}{m} Z^T Z = \frac{1}{m} (XU)^T (XU) = U^T \left(\frac{1}{m} X^T X \right) U = U^T \Sigma U$$

Since Σ is symmetric, it has real eigenvalues. Its eigen decomposition is

$$\Sigma = U \Lambda U^T$$

where

$$U = \begin{bmatrix} | & & | \\ u_1 & \dots & u_n \\ | & & | \end{bmatrix}, \Lambda = \begin{bmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_n \end{bmatrix}$$

Then

$$\text{cov}(Z) = U^T (U \Lambda U^T) U = \Lambda$$

Covariance Interpretation of PCA

PCA removes the “redundancy” (or noise) in input data X :

Let $Z = XU$ be the PCA projected data,

$$\text{cov}(Z) = \frac{1}{m} Z^T Z = \frac{1}{m} (XU)^T (XU) = U^T \left(\frac{1}{m} X^T X \right) U = U^T \Sigma U$$

Since Σ is symmetric, it has real eigenvalues. Its eigen decomposition is

$$\Sigma = U \Lambda U^T$$

where

$$U = \begin{bmatrix} | & & | \\ u_1 & \dots & u_n \\ | & & | \end{bmatrix}, \Lambda = \begin{bmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_n \end{bmatrix}$$

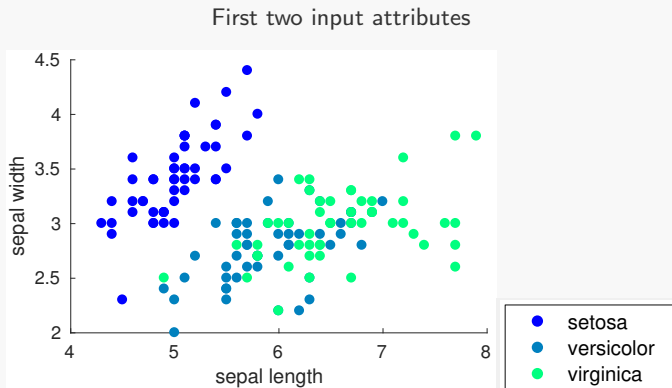
Then

$$\text{cov}(Z) = U^T (U \Lambda U^T) U = \Lambda$$

The principal component transformation XU diagonalizes the sample covariance matrix of X

PCA Example: Iris Dataset

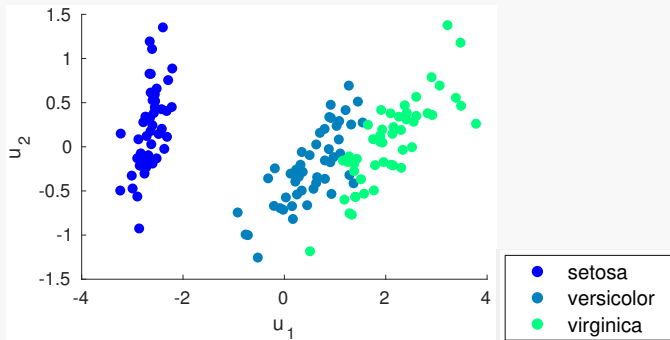
- ▶ 150 samples
- ▶ input feature dimension: 4



PCA Example: Iris Dataset

- ▶ 150 samples
- ▶ input feature dimension: 4

PCA Projection on 2 Principal Components



% of variance explained by PC1: 73%, by PC2: 22%

PCA Example: Eigenfaces

Learning image representations for face recognition using PCA [Turk and Pentland CVPR 1991]

Training data

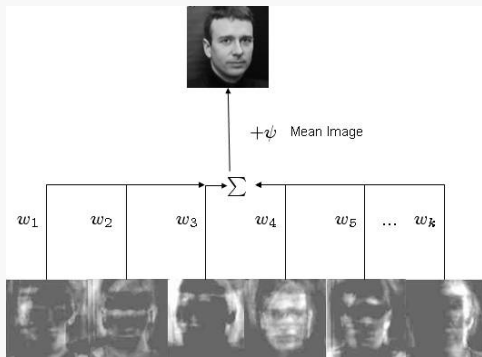


Eigenfaces: k principal components



PCA Example: Eigenfaces

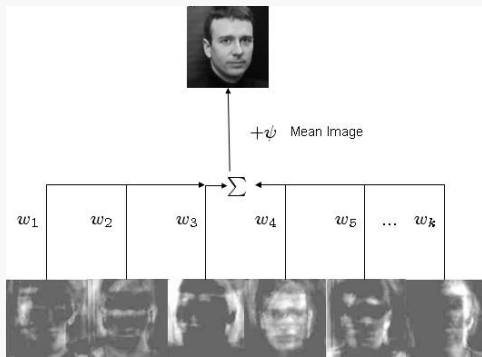
Each face image is a linear combination of the **eigenfaces** (principal components)



Each image is represented by k weights

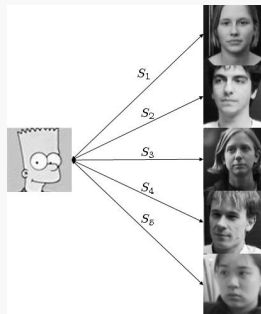
PCA Example: Eigenfaces

Each face image is a linear combination of the **eigenfaces** (principal components)



Each image is represented by k weights

Recognize faces by classifying the weight vectors. e.g. k-Nearest Neighbor



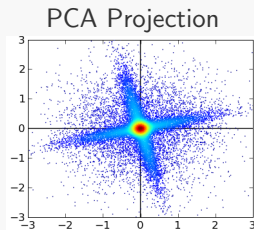
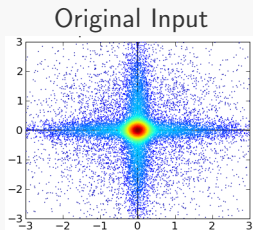
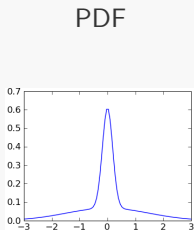
PCA Limitations

- ▶ Assumes input data is real and continuous
- ▶ Assumes **approximate normality** of input space (but may still work well on non-normally distributed data in practice) ← *sample mean & covariance must be sufficient statistics*

PCA Limitations

- ▶ Assumes input data is real and continuous
- ▶ Assumes **approximate normality** of input space (but may still work well on non-normally distributed data in practice) ← *sample mean & covariance must be sufficient statistics*

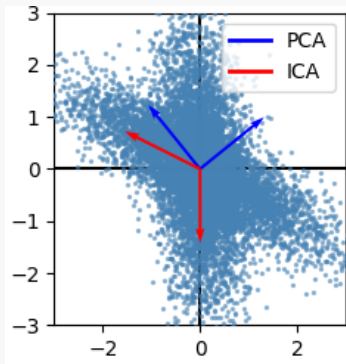
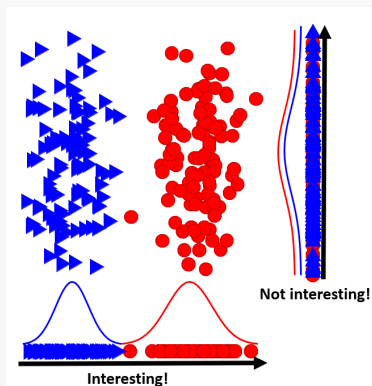
Example of strongly non-normal distributed input:



PCA Limitations

PCA results may not be useful when

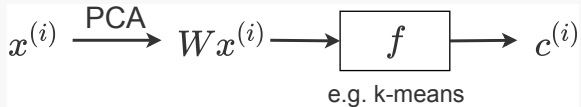
- ▶ Axes of larger variance is less 'interesting' than smaller ones.
- ▶ Axes of variations are not orthogonal;
- ▶ Data has non-linear relationships (see kernel PCA)



Kernel PCA

Kernel PCA

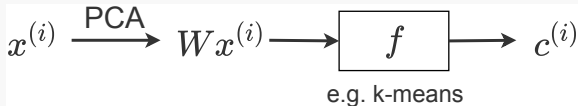
Feature extraction using PCA



Linear PCA assumes data are separable in \mathbb{R}^n

Kernel PCA

Feature extraction using PCA



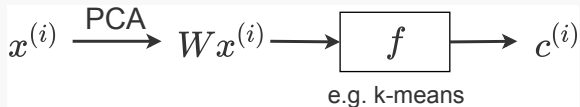
Linear PCA assumes data are separable in \mathbb{R}^n

A non-linear generalization

- ▶ Project data into higher dimension using feature mapping $\phi: \mathbb{R}^n \rightarrow \mathbb{R}^d$ ($d \geq n$)

Kernel PCA

Feature extraction using PCA



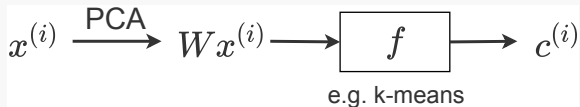
Linear PCA assumes data are separable in \mathbb{R}^n

A non-linear generalization

- ▶ Project data into higher dimension using feature mapping $\phi: \mathbb{R}^n \rightarrow \mathbb{R}^d$ ($d \geq n$)
- ▶ Feature mapping is defined by a kernel function $K(x^{(i)}, x^{(j)}) = \phi(x^{(i)})^T \phi(x^{(j)})$ or kernel matrix $K \in \mathbb{R}^{m \times m}$

Kernel PCA

Feature extraction using PCA



Linear PCA assumes data are separable in \mathbb{R}^n

A non-linear generalization

- ▶ Project data into higher dimension using feature mapping $\phi: \mathbb{R}^n \rightarrow \mathbb{R}^d$ ($d \geq n$)
- ▶ Feature mapping is defined by a kernel function $K(x^{(i)}, x^{(j)}) = \phi(x^{(i)})^T \phi(x^{(j)})$ or kernel matrix $K \in \mathbb{R}^{m \times m}$
- ▶ We can now perform standard PCA in the feature space

Kernel PCA

(Bernhard Schoelkopf, Alexander J. Smola, and Klaus-Robert Mueller. 1999. *Kernel principal component analysis*. In *Advances in kernel methods*) Sample covariance matrix of feature mapped data (assuming $\phi(x)$ is centered)

$$\Sigma = \frac{1}{m} \sum_{i=1}^m \phi(x^{(i)}) \phi(x^{(i)})^T \in \mathbb{R}^{d \times d}$$

Kernel PCA

(Bernhard Schoelkopf, Alexander J. Smola, and Klaus-Robert Mueller. 1999. *Kernel principal component analysis*. In *Advances in kernel methods*) Sample covariance matrix of feature mapped data (assuming $\phi(x)$ is centered)

$$\Sigma = \frac{1}{m} \sum_{i=1}^m \phi(x^{(i)}) \phi(x^{(i)})^T \in \mathbb{R}^{d \times d}$$

Let $(\lambda_k, u_k), k = 1, \dots, d$ be the eigen decomposition of Σ :

$$\Sigma u_k = \lambda_k u_k$$

Kernel PCA

(Bernhard Schoelkopf, Alexander J. Smola, and Klaus-Robert Mueller. 1999. *Kernel principal component analysis*. In *Advances in kernel methods*) Sample covariance matrix of feature mapped data (assuming $\phi(x)$ is centered)

$$\Sigma = \frac{1}{m} \sum_{i=1}^m \phi(x^{(i)}) \phi(x^{(i)})^T \in \mathbb{R}^{d \times d}$$

Let $(\lambda_k, u_k), k = 1, \dots, d$ be the eigen decomposition of Σ :

$$\Sigma u_k = \lambda_k u_k$$

PCA projection of $x^{(l)}$ onto the k th principal component u_k :

$$\phi(x^{(l)})^T u_k$$

Kernel PCA

(Bernhard Schoelkopf, Alexander J. Smola, and Klaus-Robert Mueller. 1999. *Kernel principal component analysis*. In *Advances in kernel methods*) Sample covariance matrix of feature mapped data (assuming $\phi(x)$ is centered)

$$\Sigma = \frac{1}{m} \sum_{i=1}^m \phi(x^{(i)}) \phi(x^{(i)})^T \in \mathbb{R}^{d \times d}$$

Let $(\lambda_k, u_k), k = 1, \dots, d$ be the eigen decomposition of Σ :

$$\Sigma u_k = \lambda_k u_k$$

PCA projection of $x^{(l)}$ onto the k th principal component u_k :

$$\phi(x^{(l)})^T u_k$$

How to avoid evaluating $\phi(x)$ explicitly?

The Kernel Trick

Represent projection $\phi(x^{(l)})^T u_k$ using kernel function K :

- Write u_k as a linear combination of $\phi(x^{(1)}), \dots, \phi(x^{(m)})$:

$$u_k = \sum_{i=1}^m \alpha_k^i \phi(x^{(i)})$$

The Kernel Trick

Represent projection $\phi(x^{(l)})^T u_k$ using kernel function K :

- Write u_k as a linear combination of $\phi(x^{(1)}), \dots, \phi(x^{(m)})$:

$$u_k = \sum_{i=1}^m \alpha_k^i \phi(x^{(i)})$$

- PCA projection of $x^{(l)}$ using kernel function K :

$$\phi(x^{(l)})^T u_k = \phi(x^{(l)})^T \sum_{i=1}^m \alpha_k^i \phi(x^{(i)}) = \sum_{i=1}^m \alpha_k^i K(x^{(l)}, x^{(i)})$$

How to find α_k^i 's directly ?

The Kernel Trick

Kth eigenvector equation:

$$\Sigma u_k = \left(\frac{1}{m} \sum_{i=1}^m \phi(x^{(i)}) \phi(x^{(i)})^T \right) u_k = \lambda_k u_k$$

- ▶ Substitute $u_k = \sum_{i=1}^m \alpha_k^{(i)} \phi(x^{(i)})$, we obtain

$$K \alpha_k = \lambda_k m \alpha_k$$

where $\alpha_k = \begin{bmatrix} \alpha_k^1 \\ \vdots \\ \alpha_k^m \end{bmatrix}$ can be solved by eigen decomposition of K

The Kernel Trick

Kth eigenvector equation:

$$\sum u_k = \left(\frac{1}{m} \sum_{i=1}^m \phi(x^{(i)}) \phi(x^{(i)})^T \right) u_k = \lambda_k u_k$$

- ▶ Substitute $u_k = \sum_{i=1}^m \alpha_k^{(i)} \phi(x^{(i)})$, we obtain

$$K \alpha_k = \lambda_k m \alpha_k$$

where $\alpha_k = \begin{bmatrix} \alpha_k^1 \\ \vdots \\ \alpha_k^m \end{bmatrix}$ can be solved by eigen decomposition of K

- ▶ Normalize α_k such that $u_k^T u_k = 1$:

$$u_k^T u_k = \sum_{i=1}^m \sum_{j=1}^m \alpha_k^i \alpha_k^j \phi(x^{(i)})^T \phi(x^{(j)}) = \alpha_k^T K \alpha_k = \lambda_k m (\alpha_k^T \alpha_k)$$

$$\|\alpha_k\|^2 = \frac{1}{\lambda_k m}$$

Kernel PCA

When $\mathbb{E}[\phi(x)] \neq 0$, we need to center $\phi(x)$:

$$\tilde{\phi}(x^{(i)}) = \phi(x^{(i)}) - \frac{1}{m} \sum_{l=1}^m \tilde{\phi}(x^{(l)})$$

Kernel PCA

When $\mathbb{E}[\phi(x)] \neq 0$, we need to center $\phi(x)$:

$$\tilde{\phi}(x^{(i)}) = \phi(x^{(i)}) - \frac{1}{m} \sum_{l=1}^m \phi(x^{(l)})$$

The “centralized” kernel matrix is

$$\tilde{K}_{i,j} = \tilde{\phi}(x^{(i)})^T \tilde{\phi}(x^{(j)})$$

In matrix notation:

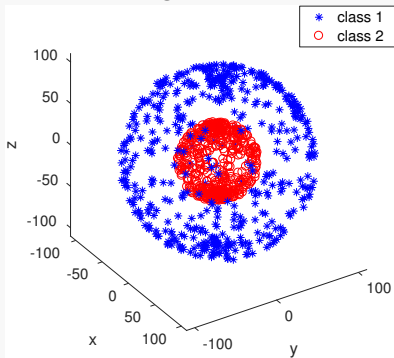
$$\tilde{K} = K - \mathbf{1}_m K - K \mathbf{1}_m + \mathbf{1}_m K \mathbf{1}_m$$

where $\mathbf{1}_m = \begin{bmatrix} 1/m & \dots & 1/m \\ \vdots & \ddots & \vdots \\ 1/m & \dots & 1/m \end{bmatrix} \in \mathbb{R}^{m \times m}$

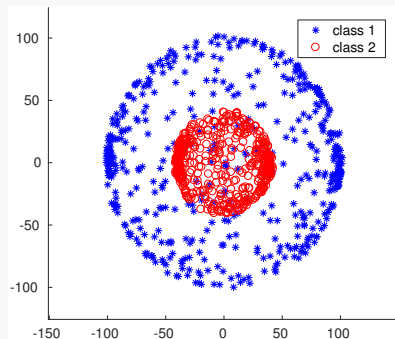
Use \tilde{K} to compute PCA

Kernel PCA Example

original data

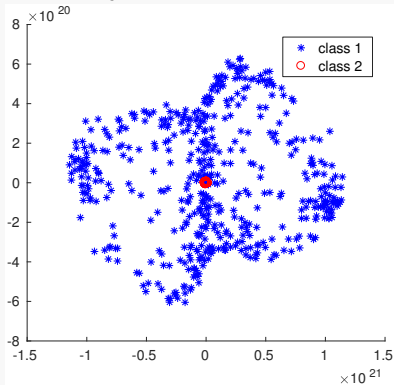


standard PCA



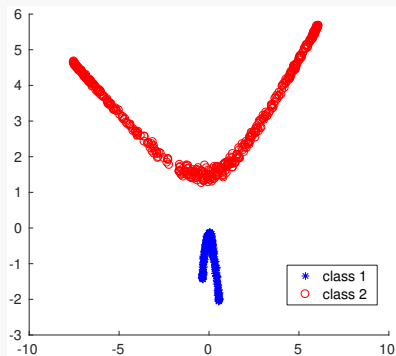
Kernel PCA Example

Polynomial kernel PCA



$$k(x, x') = (x \cdot x' + 1)^5$$

Gaussian kernel PCA



$$k(x, x') = \exp\left(-\frac{\|x-x'\|^2}{2\sigma^2}\right)$$

Discussions of kernel PCA

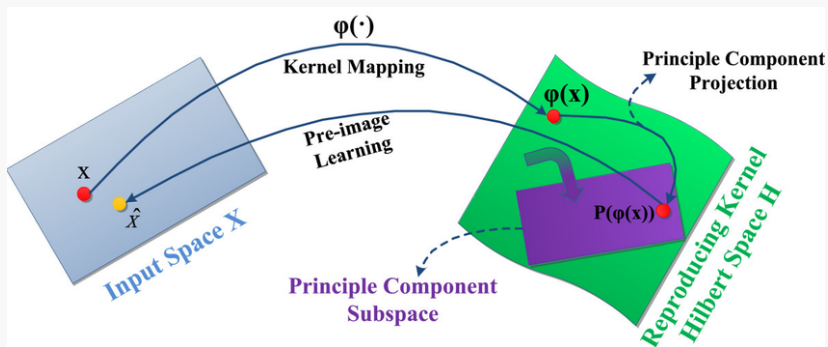
- ▶ Often used in clustering, abnormality detection, etc

Discussions of kernel PCA

- ▶ Often used in clustering, abnormality detection, etc
- ▶ Requires finding eigenvectors of $m \times m$ matrix instead of $n \times n$

Discussions of kernel PCA

- ▶ Often used in clustering, abnormality detection, etc
- ▶ Requires finding eigenvectors of $m \times m$ matrix instead of $n \times n$
- ▶ Dimension reduction by projecting to k -dimensional principal subspace is generally not possible



The Pre-Image problem: reconstruct data in input space x from feature space vectors $\phi(x)$

Summary

Representation learning

- ▶ Transform input features into “simpler” or “interpretable” representations.
- ▶ Used in feature extraction, dimension reduction, clustering etc

Summary

Representation learning

- ▶ Transform input features into “simpler” or “interpretable” representations.
- ▶ Used in feature extraction, dimension reduction, clustering etc

Unsupervised learning algorithms:

	low dimension	sparse	disentangle variations
k-means		✓	
PCA	✓		✓

Homework

Homework 3 due in one week: SVM and Neural Network

How to get partial credits for programming problems?

if result are still undesirable by the deadline,

- ▶ Write your thought process in sentences in a README file or on the problem answer sheet.
- ▶ Explain the problem you run into.