Learning from Data Lecture 9: Principal Component Analysis

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TBSI

December 3, 2021

Motivation

Unsupervised Learning (Part II): PCA

- Motivation
- ▶ Linear PCA
- Kernel PCA

Project Information: http://yangli-feasibility.com/home/

classes/lfd2021fall/project.html



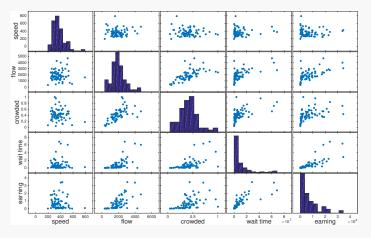
Example: Analyzing San Francisco public transit route efficiency





features	notes
speed	average speed
flow	# boarding pas-
	sengers per hour
crowded	% passenger ca-
	pacity reached
wait time	average waiting
	time at bus stop
earning	net operation rev-
	enue
:	:

Input features contain a lot of redundancy

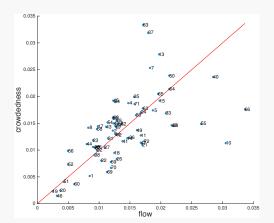


Scatter plot matrix reveals pairwise correlations among $5\ \text{major}$ features

Example of linearly dependent features

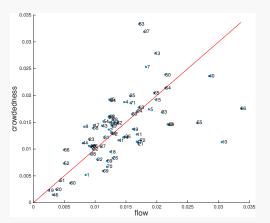
- ▶ Flow: average # boarding passengers per hour
- ► Crowdedness:

 | average # passengers on train train capacity | |



Example of linearly dependent features

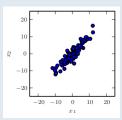
- ▶ Flow: average # boarding passengers per hour
- ► Crowdedness: average # passengers on train train capacity



How can we automatically detect and remove this redundancy?

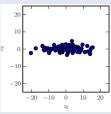
- ▶ geometric approach ← start here!
- diagonalize covariance matrix approach

- Given $\{x^{(1)}, \dots, x^{(m)}\}$, $x^{(i)} \in \mathbb{R}^n$. Find a linear, orthogonal transformation $W : \mathbb{R}^n \to \mathbb{R}^k$ of the input
 - W aligns the direction of maximum variance with the axes of the new space.



features x_1 and x_2 are strongly correlated





variations in $z = x^T W$ is mostly along the x-axis. x can be represented in 1DI

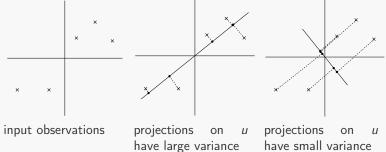
data

Direction of Maximum Variance

▶ Suppose $\mu = mean(x) = 0$, $\sigma_j = var(x_j) = 1$ (variance of jth feature)

Direction of Maximum Variance

- Suppose $\mu = mean(x) = 0$, $\sigma_j = var(x_j) = 1$ (variance of jth feature)
- Find **major axis of variation** unit vector *u*:



Learning From Data

u maximizes the variance of the projections

Motivation

Linear PCA

Kernel PCA

Principal Component Analysis (PCA)

Pearson, K. (1901), Hotelling, H. (1933) "Analysis of a complex of statistical variables into principal components". Journal of Educational Psychology.

PCA goals

- Find principal components u_1, \ldots, u_n that are mutually orthogonal (uncorrelated)
- Most of the variation in x will be accounted for by k principal components where $k \ll n$.

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- Find principal components u_1, \ldots, u_n that are mutually orthogonal (uncorrelated)
- ▶ Most of the variation in x will be accounted for by k principal components where $k \ll n$.

Main steps of (full) PCA:

- **1.** Standardize x such that Mean(x) = 0, $Var(x_i) = 1$ for all j
- 2. Find projection of x, $u_1^T x$ with maximum variance
- 3. For j = 2, ..., n, Find another projection of x, $u_i^T x$ with maximum variance, where u_i is orthogonal to u_1, \ldots, u_{i-1}

Step 1: Standardize data

Normalize x such that Mean(x) = 0 and $Var(x_j) = 1$

$$x^{(i)} := x^{(i)} - \mu \leftarrow \text{recenter}$$

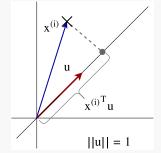
 $x_j^{(i)} := x_j^{(i)} / \sigma_j \leftarrow \text{scale by } stdev(x_j)$

Check:

$$var\left(\frac{x_j}{\sigma_j}\right) = \frac{1}{m} \sum_{i=1}^m \left(\frac{x_j^{(i)} - \mu_j}{\sigma_j}\right)^2 = \frac{1}{\sigma_j^2} \frac{1}{m} \sum_{i=1}^m \left(x_j^{(i)} - \mu_j\right)^2$$
$$= \frac{1}{\sigma_j^2} \sigma_j^2 = 1$$

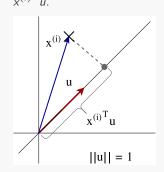
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Since ||u|| = 1, the length of $x^{(i)}$'s projection on u is $x^{(i)}^T u$.



Step 2: Find Projection with Maximum Variance

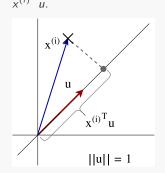
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Variance of the projections:

$$\frac{1}{m} \sum_{i=1}^{m} (x^{(i)}^{T} u - \mathbf{0})^{2} = \frac{1}{m} \sum_{i=1}^{m} u^{T} x^{(i)} x^{(i)}^{T} u$$

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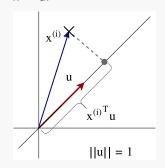


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$$= u^{T} \left(\frac{1}{m} \sum_{i=1}^{m} x^{(i)} x^{(i)}^{T} \right) u$$

$$= u^{T} \Sigma u$$

 Σ : the sample covariance matrix of $x^{(1)} \dots x^{(m)}$.

1st Principal Component

Find unit vector u_1 that maximizes variance of projections:

$$u_1 = \underset{u: \|u\| = 1}{\operatorname{argmax}} \ u^T \Sigma u \tag{1}$$

 u_1 is the **1st principal component** of X

 u_1 can be solved using optimization tools, but it has a more efficient solution:

Proposition 1

 u_1 is the largest eigenvector of covariance matrix Σ

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Proof.

 \textit{u}_1 is the largest eigenvector of covariance matrix Σ

 ${\it Proof.} \ \ {\it Generalized Lagrange function of Problem} \ \ \underline{1};$

$$L(u) = -u^T \Sigma u + \beta (u^T u - 1)$$

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Proof. Generalized Lagrange function of Problem 1:

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To minimize L(u),

$$\frac{\delta L}{\delta u} = -2\Sigma u + 2\beta u = 0 \implies \Sigma u = \beta u$$

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Therefore u_1 must be an eigenvector of Σ .

Let $u_1 = v_j$, the eigenvector with the jth largest eigenvalue λ_j ,

$$u_1^T \Sigma u_1 = v_i^T \Sigma v_j = \lambda_j v_j^T v_j = \lambda_j.$$

Hence $u_1 = v_1$, the eigenvector with the largest eigenvalue λ_1 .

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Linear PCA

Proposition 2

The jth principal component of X , u_j is the jth largest eigenvector of Σ .

Proof.

Motivation

The jth principal component of X, u_i is the jth largest eigenvector of Σ.

Proof. Consider the case j = 2,

$$u_2 = \underset{u:||u||=1, u_1^T u=0}{\operatorname{argmax}} u^T \Sigma u$$
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The Lagrangian function:

$$L(u) = -u^T \Sigma u + \beta_1 (u^T u - 1) + \beta_2 (u_1^T u)$$

Minimizing L(u) yields:

$$\beta_2=0, \Sigma u=\beta_1 u$$

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To maximize $u^T \Sigma u = \lambda$, u_2 must be the eigenvector with the second largest eigenvalue $\beta_1 = \lambda_2$. The same argument can be generalized to cases j > 2. (Use induction to prove for $j = 1 \dots n$)

We can solve PCA by solving an eigenvalue problem! Main steps of (full) PCA:

- 1. Standardize x such that Mean(x) = 0, $Var(x_i) = 1$ for all j
- 2. Compute $\Sigma = cov(x)$
- 3. Find principal components u_1, \ldots, u_n by eigenvalue decomposition: $\Sigma = U \wedge U^T$. $\leftarrow U$ is an orthogonal basis in \mathbb{R}^n

Next we project data vectors x to this new basis, which spans the **principal component space**.

▶ Projection of sample
$$x \in \mathbb{R}^n$$
 in the principal component space:

$$\underline{z}^{(i)} = \begin{bmatrix} \underline{x}^{(i)}^T \underline{u}_1 \\ \vdots \\ \underline{x}^{(i)}^T \underline{u}_n \end{bmatrix} \in \mathbb{R}^n$$

$$Z_j^{(i)} = x^{(i)}^T \underline{u}_j$$

▶ Projection of sample $x \in \mathbb{R}^n$ in the principal component space:

Linear PCA

$$z^{(i)} = \begin{bmatrix} x^{(i)} & u_1 \\ \vdots \\ x^{(i)} & u_n \end{bmatrix} \in \mathbb{R}^n$$

Matrix notation:

$$z^{(i)} = \begin{bmatrix} | & | & | \\ | u_1 & \dots & | \\ | & | \end{bmatrix}^T \underbrace{x^{(i)}}_{x^{(i)}} = \underbrace{U^T x^{(i)}}_{x^{(i)}}, \text{ or } Z = XU$$

$$\begin{bmatrix} -x^{(i)} \\ | & | \\ | & | \end{bmatrix} \begin{bmatrix} | & | \\ | & | \\ | & | \\ | & | \end{bmatrix}$$

$$\underbrace{(m \times h)}_{x^{(i)}} = \underbrace{(m \times h)}_{x^{(i)}}_{x^{(i)}} = \underbrace{(m \times h)}_{x^{(i)}} = \underbrace{(m \times h)}_{x^{(i)}}_{x^{(i)}} = \underbrace{(m \times h)}_{x^{(i)}} = \underbrace{(m \times h)$$

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The truncated transformation $Z_k = XU_k$ keeping only the first k principal components is used for **dimension reduction**.

Motivation

The variance of principal component projections are

$$Var(x^Tu_j) = u_j^T \Sigma u_j = \lambda_j \text{ for } j = 1, \dots, n$$

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$$\sum_{i=1}^n Var(x^T u_i)$$

• % of variance explained by the jth principal component:

i.e. projections are uncorrelated

The variance of principal component projections are

$$Var(x^Tu_j) = u_j^T \Sigma u_j = \lambda_j \text{ for } j = 1, \dots, n$$

- % of variance explained by the *j*th principal component: $\frac{(\lambda_j)}{\sum_{i=1}^{n} \lambda_i}$ i.e. projections are uncorrelated
- % of variance accounted for by retaining the first k principal

components
$$(\underline{k} \le n)$$
: $\frac{\sum_{j=1}^{k} \lambda_j}{\sum_{j=1}^{n} \lambda_j}$

Another geometric interpretation of <u>PCA</u> is minimizing projection residuals. (see homework!)

Covariance Interpretation of PCA

covariance diagnalization cov(x)= + XX

PCA removes the "redundancy" (or noise) in input data X: Let Z = XU be the PCA projected data.

$$cov(Z) = \frac{1}{m}Z^{T}Z = \frac{1}{m}(\underbrace{XU})^{T}(\underbrace{XU}) = U^{T}(\underbrace{\frac{1}{m}X^{T}X})U = \underline{U^{T}\Sigma U}$$

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Since Σ is symmetric, it has real eigenvalues. Its eigen decomposition is

$$\Sigma = U \Lambda U^T$$

where

$$U = \begin{bmatrix} | & & | \\ u_1 & \dots & u_n \\ | & & | \end{bmatrix}, \Lambda = \begin{bmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_n \end{bmatrix}$$

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Then

 $cov(Z) = U^{T}(U \wedge U^{T})U = \Lambda^{C}$ diagonal metrix (

Learning From Data

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The principal component transformation XU diagonalizes the sample covariance matrix of X

PCA Dimension reduction

- Find principal components $\underline{u_1}, \dots, \underline{u_n}$ that are mutually orthogonal (uncorrelated)
- Most of the variations in x will be accounted for by k principal components where $k \ll n$.

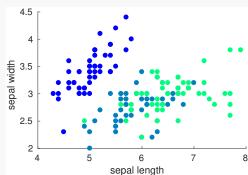
Main steps

- 1. Standardize \underline{x} such that Mean(x) = 0, $Var(x_j) = 1$ for all j
- 2. Compute $\Sigma = cov(x)$
- **3.** Find principal components u_1, \ldots, u_n by eigenvalue decomposition: $\Sigma = U \wedge U^T$. $\leftarrow U$ is an orthogonal basis in \mathbb{R}^n
- **4.** Project data on first the k principal components: $z = [x^T u_1, \dots, x^T u_k]^T$

PCA Example: Iris Dataset

- ▶ 150 samples
- ▶ input feature dimension: 4



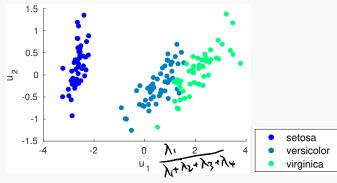


setosa versicolor virginica

PCA Example: Iris Dataset

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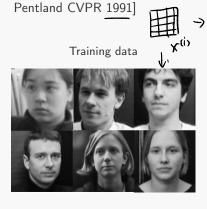




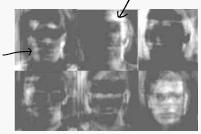
% of variance explained by PC1: 73%, by PC2: 22% = 95%

PCA Example: Eigenfaces

Learning image representations for face recognition using PCA [Turk and

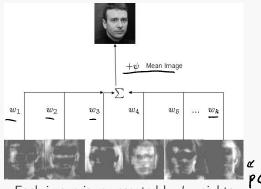


Eigenfaces: k principal components



PCA Example: Eigenfaces

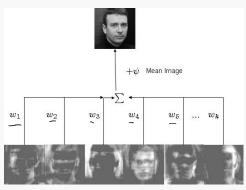
Each face image is a linear combination of the eigenfaces (principal components)



Each image is represented by k weights

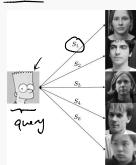
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Recognize faces by classifying the weight vectors. e.g. k-Nearest Neighbor

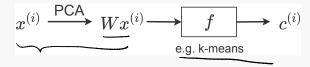


Linear PCA

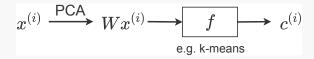
Kernel PCA

Kernel PCA

Motivation



Linear PCA assumes data are separable in \mathbb{R}^n



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A non-linear generalization

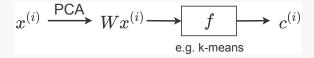
Project data into higher dimension using feature mapping $\phi: \mathbb{R}^n \to \mathbb{R}^d \ (d \ge n)$

$$x^{(i)} \xrightarrow{\operatorname{PCA}} Wx^{(i)} \xrightarrow{\qquad \qquad} f \xrightarrow{\qquad \qquad} c^{(i)}$$
 e.g. k-means

Linear PCA assumes data are separable in \mathbb{R}^n

A non-linear generalization

- ▶ Project data into higher dimension using feature mapping $\phi : \mathbb{R}^n \to \mathbb{R}^d \ (d \ge n)$
- Feature mapping is defined by a kernel function $K\left(x^{(i)},x^{(j)}\right) = \phi(x^{(i)})^T\phi(x^{(j)})$ or kernel matrix $K \in \mathbb{R}^{m \times m}$



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A non-linear generalization

- ▶ Project data into higher dimension using feature mapping $\phi : \mathbb{R}^n \to \mathbb{R}^d \ (d \ge n)$
- Feature mapping is defined by a kernel function $K(x^{(i)}, x^{(j)}) = \phi(x^{(i)})^T \phi(x^{(j)})$ or kernel matrix $K \in \mathbb{R}^{m \times m}$
- We can now perform standard PCA in the feature space

(Bernhard Schoelkopf, Alexander J. Smola, and Klaus-Robert Mueller. 1999. Kernel principal component analysis. In Advances in kernel methods) Sample covariance matrix of feature mapped data (assuming $\phi(x)$ is centered)

$$\sum = \frac{1}{m} \sum_{i=1}^{m} \phi(x^{(i)}) \phi(x^{(i)})^{T} \in \mathbb{R}^{\frac{d \times d}{d}}$$

$$\chi = \begin{bmatrix} -x^{(i)} \\ -x^{(m)} \end{bmatrix} \oint \begin{bmatrix} \phi(x^{(i)}) - \\ \phi(x^{(m)}) \end{bmatrix}$$

(Bernhard Schoelkopf, Alexander J. Smola, and Klaus-Robert Mueller. 1999. *Kernel principal component analysis*. In Advances in kernel methods) Sample covariance matrix of feature mapped data (assuming $\phi(x)$ is centered)

$$\Sigma = \frac{1}{m} \sum_{i=1}^{m} \phi(x^{(i)}) \phi(x^{(i)})^{T} \in \mathbb{R}^{d \times d}$$

Let $(\lambda_k, u_k), k = 1, ..., d$ be the eigen decomposition of Σ :

$$\sum u_k = \lambda_k u_k$$

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PCA projection of $x^{(1)}$ onto the *kth* principal component u_k :

$$\phi(x^{(I)})^T u_k$$

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PCA projection of $x^{(I)}$ onto the kth principal component u_k :

$$\phi(x^{(I)})^T u_k$$

How to avoid evaluating $\phi(x)$ explicitly?

The Kernel Trick

Represent projection $\phi(x^{(l)})^T u_k$ using kernel function K:

• Write u_k as a linear combination of $\phi(x^{(1)}), \ldots, \phi(x^{(m)})$:

$$\underline{u}_{k} = \sum_{i=1}^{m} \alpha_{k}^{i} \underline{\phi(x^{(i)})}$$

$$\sum_{i=1}^{m} \omega_{k} = \lambda_{k} u_{k}$$

$$\left(\frac{1}{m} \sum_{i=1}^{m} \phi(x^{(i)}) \phi(x^{(i)})^{T}\right) u_{k} = \lambda_{k} u_{k}$$

$$\frac{1}{m \lambda_{k}} \sum_{i=1}^{m} \phi(x^{(i)}) \phi(x^{(i)})^{T} u_{k} = u_{k}$$

$$\sum_{i=1}^{m} \frac{1}{m \lambda_{k}} \phi(x^{(i)})^{T} u_{k}$$

$$\frac{1}{m \lambda_{k}} \sum_{i=1}^{m} \psi(x^{(i)}) \phi(x^{(i)})^{T} u_{k}$$

Represent projection $\phi(x^{(l)})^T u_k$ using kernel function K:

• Write u_k as a linear combination of $\phi(x^{(1)}), \ldots, \phi(x^{(m)})$:

$$u_k = \sum_{i=1}^m \underline{\alpha}_k^i \phi(x^{(i)})$$

▶ PCA projection of $x^{(l)}$ using kernel function K:

$$\phi(x^{(l)})^T \underline{u_k} = \phi(x^{(l)})^T \sum_{i=1}^m \alpha_k^i \phi(x^{(i)}) = \sum_{i=1}^m \alpha_k^i \underbrace{K(x^{(l)}, x^{(i)})}$$

How to find α_k^i 's directly ?

Kth eigenvector equation:

$$\underline{\sum u_k} = \left(\underline{\frac{1}{m}} \sum_{i=1}^m \phi(x^{(i)}) \phi(x^{(i)})^T\right) \underline{u_k} = \lambda_k u_k$$

Substitute $u_k = \sum_{i=1}^m (\alpha_k^{(i)}) \phi(x^{(i)})$, we obtain

$$K\alpha_{k} = \lambda_{k} m \alpha_{k}$$

$$F K \omega_{k}$$

$$Cm \times m (m \cdot 1)$$

$$G K \omega_{k}$$

$$G K \omega_{$$

$$(m \times m) (m \times 1)$$

The Kernel Trick

Kth eigenvector equation:

$$\sum u_k = \left(\frac{1}{m} \sum_{i=1}^m \phi(x^{(i)}) \phi(x^{(i)})^T\right) u_k = \lambda_k u_k$$

Substitute $u_k = \sum_{i=1}^m \alpha_k^{(i)} \phi(x^{(i)})$, we obtain

$$K\alpha_k = \lambda_k m\alpha_k$$

where $\alpha_k = \begin{bmatrix} \alpha_k^* \\ \vdots \\ m \end{bmatrix}$ can be solved by eigen decomposition of K

Normalize $\underline{\alpha}_k$ such that $u_k^T u_k = 1$:

$$u_k^T u_k = \sum_{i=1}^m \sum_{j=1}^m \alpha_k^i \alpha_k^j \phi(x^{(i)})^T \phi(x^{(j)}) = \alpha_k^T K \alpha_k = \lambda_k m \alpha_k^T \alpha_k = 1$$

$$\|\alpha_k\|^2 = \frac{1}{\lambda_k m}$$

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Motivation

When $\mathbb{E}[\phi(x)] \neq 0$, we need to center $\phi(x)$:

$$\widetilde{\phi}(x^{(i)}) = \phi(x^{(i)}) - \frac{1}{m} \sum_{l=1}^{m} \widetilde{\phi}(x^{(l)})$$

Kernel PCA

When $\mathbb{E}[\phi(x)] \neq 0$, we need to center $\phi(x)$:

$$\widetilde{\phi}(x^{(i)}) = \phi(x^{(i)}) - \frac{1}{m} \sum_{l=1}^{m} \widetilde{\phi}(x^{(l)})$$

The "centralized" kernel matrix is

$$\widetilde{K}_{i,j} = \widetilde{\phi}(x^{(i)})^T \widetilde{\phi}(x^{(j)})$$

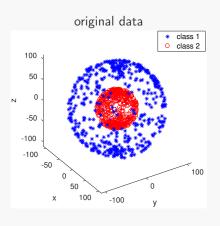
In matrix notation:

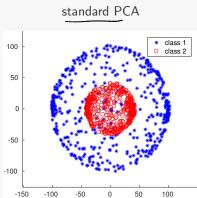
$$\widetilde{K} = K - \frac{1}{1m}K - K\mathbf{1}_m + \mathbf{1}_m K\mathbf{1}_m$$

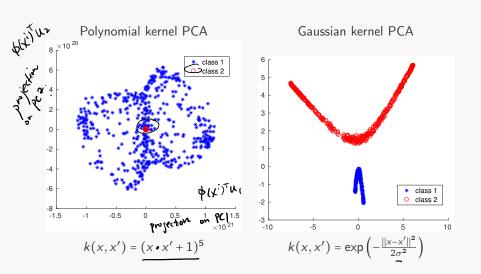
where
$$\underline{\mathbf{1}_m} = \begin{bmatrix} 1/m & \dots & 1/m \\ \vdots & \ddots & \vdots \\ 1/m & \dots & 1/m \end{bmatrix} \in \mathbb{R}^{m \times m}$$

Use \widetilde{K} to compute PCA

Kernel PCA Example







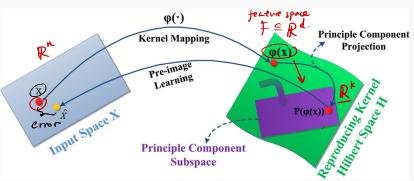
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▶ Often used in clustering, abnormality detection, etc

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- Often used in clustering, abnormality detection, etc.
- Requires finding eigenvectors of $m \times m$ matrix instead of $n \times n$
- Dimension reduction by projecting to k-dimensional principal subspace is generally not possible



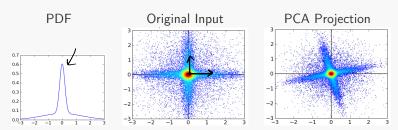
The Pre-Image problem: reconstruct data in input space x from feature space vectors $\phi(x)$

PCA Limitations

- Assumes input data is real and continuous
- Assumes approximate normality of input space (but may still work well on non-normally distributed data in practice) ← sample mean & covariance must be sufficient statistics

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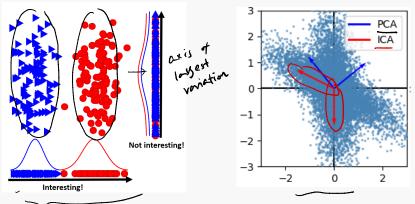
Example of strongly non-normal distributed input:



A Lillitations

PCA results may not be useful when

- Axes of larger variance is less 'interesting' than smaller ones.
 - Axes of variations are not orthogonal;



Representation learning

- Transform input features into "simpler" or "interpretable" representations.
- ▶ Used in feature extraction, dimension reduction, clustering etc

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Unsupervised learning algorithms:

