

Learning From Data

Lecture 2: Linear Regression & Logistic Regression

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Today's Lecture

Supervised Learning (Part I)

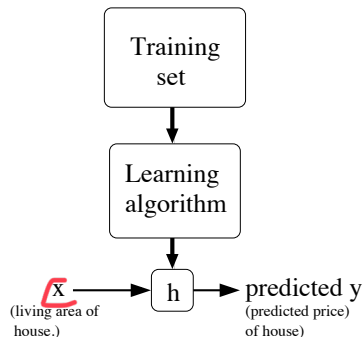
- ▶ Linear Regression
- ▶ Binary Classification
- ▶ Multi-Class Classification

Review: Supervised Learning

- ▶ Input space: \mathcal{X} , Target space: \mathcal{Y}

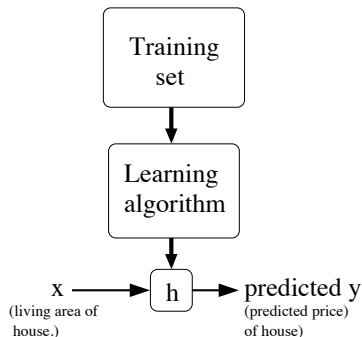
Review: Supervised Learning

- ▶ Input space: \mathcal{X} , Target space: \mathcal{Y}
- ▶ Given training examples, we want to learn a **hypothesis** function $h : \mathcal{X} \rightarrow \mathcal{Y}$ so that $h(x)$ is a "good" predictor for the corresponding y .



Review: Supervised Learning

- ▶ Input space: \mathcal{X} , Target space: \mathcal{Y}
- ▶ Given training examples, we want to learn a **hypothesis** function $h : \mathcal{X} \rightarrow \mathcal{Y}$ so that $h(x)$ is a "good" predictor for the corresponding y .



- ▶ y is discrete (categorical): **classification problem**
- ▶ y is continuous (real value): **regression problem**

Review: Inference vs Learning

Given training data of x and y ,

Inference

knowing the structure of f , find good models to describe f . i.e. model the data generation process ← focus of statistics

Prediction

given **future** data samples of x , predict the corresponding output data y using f . ← focus of machine learning

Linear Regression

Linear Regression Model

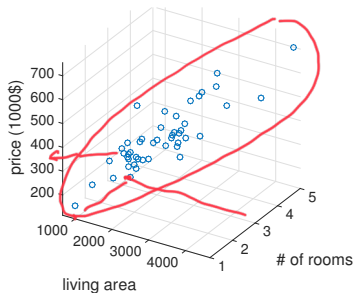
Ordinary Least Square

Maximum Likelihood Estimation

Linear Regression

Example: predict Portland housing price

<u>Living area (ft^2)</u>	<u># bedrooms</u>	<u>Price (\$1000)</u>
x_1	x_2	y
2104	3	400
1600	3	330
2400	3	369
\vdots	\vdots	\vdots



Linear Approximation

A linear model

$$h(x) = \theta_0 + \theta_1 x_1 + \theta_2 x_2$$

θ_i 's are called **parameters**.

Linear Approximation

A linear model

$$\underline{h(x)} = \theta_0 + \theta_1 x_1 + \theta_2 x_2$$

θ_i 's are called **parameters**.

Using vector notation,

$$h(x) = \theta^T x, \quad \text{where } \underline{\theta} = \begin{bmatrix} \theta_0 \\ \theta_1 \\ \theta_2 \end{bmatrix}, \quad \underline{x} = \begin{bmatrix} 1 \\ x_1 \\ x_2 \end{bmatrix}$$

Alternative Notation

$$h(x) = \overset{\theta_1}{w_1}x_1 + \overset{\theta_2}{w_2}x_2 + \overset{\theta_0}{b}$$

w_1, w_2 are called **weights**, b is called the **bias**

$$h(x) = \underline{w^T x} + b, \quad \text{where } w = \begin{bmatrix} w_1 \\ w_2 \end{bmatrix}, \quad x = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}$$

$b \in \mathbb{R}$

Apply model to new data

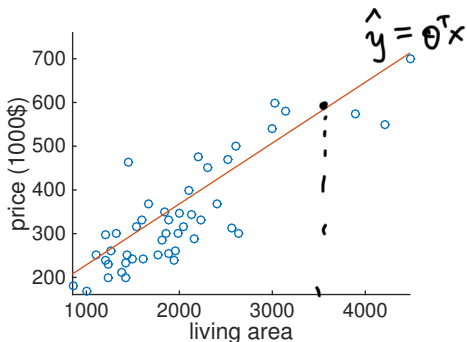
Suppose we have the optimal parameters θ , e.g.

```
> h = LinearRegression().fit(X, y)
> theta = h.coef
array([89.60, 0.1392, -8.738])
```

θ_0 θ_1 θ_2

make a prediction of new feature x :

$$\hat{y} = h_{\theta}(x) = \theta^T x$$



Model Estimation

How to estimate model parameters θ (or w and b) from data?

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Minimize sum of the prediction error squared (least square error)
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Maximum Likelihood Estimation

- ▶ Assume the data are generated from $h(x)$ with some noise distribution.
- ▶ Determines the parameters θ most likely to produce the observed data.

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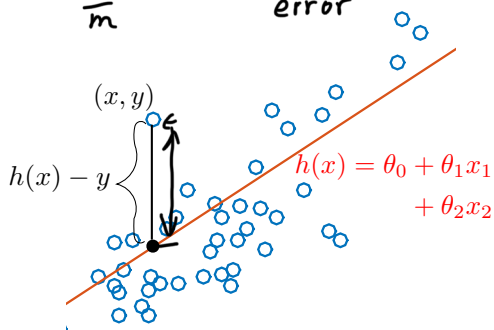
Other estimation methods exist, e.g. Bayesian estimation

Ordinary Least Square

Cost function:

$$J(\theta) = \frac{1}{2m} \sum_{i=1}^m (h(x^{(i)}) - y^{(i)})^2$$

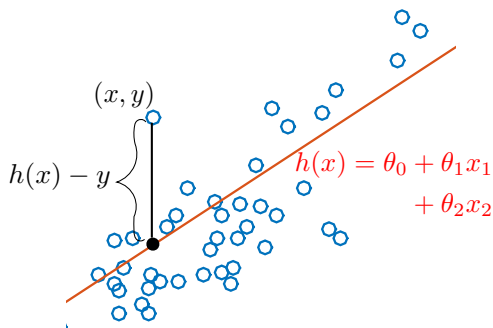
prediction / ground truth
error



Ordinary Least Square

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- ▶ This model is called **ordinary least square**

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Ordinary Least square problem

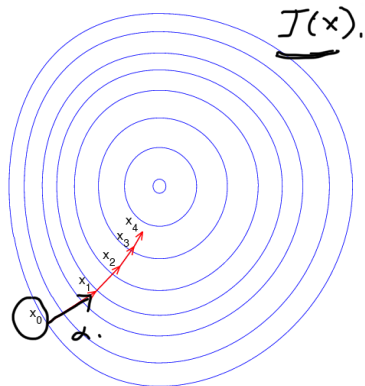
$$\begin{aligned} & \min_{\theta} J(\theta) \\ &= \min_{\theta} \frac{1}{2} \sum_{i=1}^m (h(x^{(i)}) - y^{(i)})^2 \end{aligned}$$

How to minimize $J(\theta)$?

- ▶ Numerical solution: gradient descent, Newton's method
- ▶ Analytical solution: normal equation

Gradient descent

A first-order iterative optimization algorithm for finding the minimum of a function $J(\theta)$.



Key idea

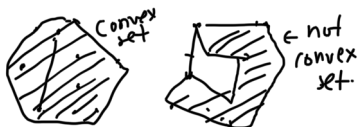
Start at an initial guess, repeatedly change θ to decrease $J(\theta)$:

$$\theta := \theta - \alpha \nabla J(\theta)$$

↑

α is the **learning rate**

Review: Convex function

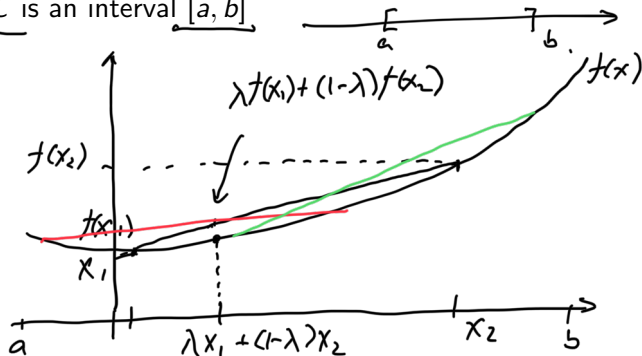


Definition

A function $f(x)$ is **convex** on a convex set C if for any $x_1, x_2 \in C$ and $0 \leq \lambda \leq 1$,

$$\underline{f(\lambda x_1 + (1 - \lambda)x_2)} \leq \underline{\lambda f(x_1) + (1 - \lambda)f(x_2)}$$

e.g. C is an interval $[a, b]$



Review: Convex function

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A function $f(x)$ is **convex** on a convex set C if for any $x_1, x_2 \in C$ and $0 \leq \lambda \leq 1$,

$$f(\lambda x_1 + (1 - \lambda)x_2) \leq \lambda f(x_1) + (1 - \lambda)f(x_2)$$

e.g. C is an interval $[a, b]$

Theorem

If $J(\theta)$ is convex, gradient descent finds the global minimum.

For the ordinary least square problem,

$$J(\theta) = \frac{1}{2} \sum_{i=1}^m \underbrace{(h(x^{(i)}) - y^{(i)})^2}_{(\theta^T x^{(i)} - y^{(i)})^2} = \frac{1}{2} \sum_{i=1}^m (\theta^T x^{(i)} - y^{(i)})^2,$$

$$\underbrace{\nabla J(\theta)}_{\theta = \begin{bmatrix} \theta_1 \\ \vdots \\ \theta_n \end{bmatrix}} = \begin{bmatrix} \frac{\partial J(\theta)}{\partial \theta_1} \\ \vdots \\ \frac{\partial J(\theta)}{\partial \theta_n} \end{bmatrix}, \text{ where } \frac{\partial J(\theta)}{\partial \theta_j} = \frac{\partial}{\partial \theta_j} \left(\frac{1}{2} \sum_{i=1}^m (\theta^T x^{(i)} - y^{(i)})^2 \right)$$

$$= \frac{1}{2} \sum_{i=1}^m 2(\theta^T x^{(i)} - y^{(i)}) \underbrace{\frac{\partial}{\partial \theta_j} (\theta^T x^{(i)} - y^{(i)})}_{x_j^i}$$

$$= \sum_{i=1}^m x_j^i (\theta^T x^{(i)} - y^{(i)})$$

For the ordinary least square problem,

$$J(\theta) = \frac{1}{2} \sum_{i=1}^m (h(x^{(i)}) - y^{(i)})^2 = \frac{1}{2} \sum_{i=1}^m (\theta^T x^{(i)} - y^{(i)})^2,$$

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$$= \sum_{i=1}^m (\theta^T x^{(i)} - y^{(i)}) x_j^{(i)}$$

Gradient descent for ordinary least square

Gradient of cost function: $\nabla J(\theta) = \sum_{i=1}^m (\theta^T x^{(i)} - y^{(i)}) x_j^{(i)}$
Gradient descent update: $\theta := \theta - \alpha \nabla J(\theta)$

$\underbrace{\hspace{10em}}_{-(y^{(i)} - \theta^T x^{(i)})}$

Batch Gradient Descent

Repeat until convergence {
 $\theta_j = \theta_j \oplus \alpha \underbrace{\sum_{i=1}^m (y^{(i)} - h_{\theta}(x^{(i)})) x_j^{(i)}}_{\nabla J(\theta)}$ for every j
}

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```
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   $\theta_j = \theta_j + \alpha \sum_{i=1}^m (y^{(i)} - h_{\theta}(x^{(i)})) x_j^{(i)}$  for every j  
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```

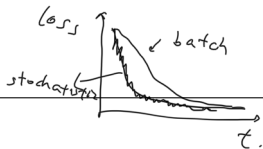
θ is only updated after we have seen all m training samples.

Batch gradient descent

Repeat until convergence {
 $\theta_j = \theta_j + \alpha \sum_{i=1}^m (y^{(i)} - h_{\theta}(x^{(i)})) x_j^{(i)}$ for every j
}

Stochastic gradient descent

Repeat until convergence {
for $i = 1 \dots m$ {
 $\theta_j = \theta_j + \alpha (y^{(i)} - h_{\theta}(x^{(i)})) x_j^{(i)}$ for every j
}
}



θ is updated each time a training example is read

$$\frac{1}{n} (X^T X)^{-1} X^T y$$

Handwritten notes: λI is written below the denominator, and a circled $+ \lambda$ is written below the $X^T X$ term, indicating the addition of a regularization term to the inverse matrix.

Batch gradient descent

```
Repeat until convergence{  
   $\theta_j = \theta_j + \alpha \sum_{i=1}^m (y^{(i)} - h_{\theta}(x^{(i)}))x_j^{(i)}$  for every j  
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Stochastic gradient descent

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Repeat until convergence{  
  for  $i = 1 \dots m$  {  
     $\theta_j = \theta_j + \alpha (y^{(i)} - h_{\theta}(x^{(i)}))x_j^{(i)}$  for every j  
  }  
}
```

θ is updated each time a training example is read

- ▶ Stochastic gradient descent gets θ close to minimum much faster
- ▶ Good for regression on large data

Minimize $J(\theta)$ Analytically

$$x^{(i)} = \begin{bmatrix} x_1^{(i)} \\ \vdots \\ x_n^{(i)} \end{bmatrix}$$

The matrix notation

$$\underbrace{X}_{(m \times n)} = \begin{bmatrix} - (x^{(1)})^T - \\ - (x^{(2)})^T - \\ \vdots \\ - (x^{(m)})^T - \end{bmatrix}, \quad \vec{y} = \begin{bmatrix} y^{(1)} \\ y^{(2)} \\ \vdots \\ y^{(m)} \end{bmatrix} \quad (m \times 1)$$

X is called the **design matrix**.

Minimize $J(\theta)$ Analytically

The matrix notation

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X is called the **design matrix**. The least square function can be written as

$$J(\theta) = \frac{1}{2} \underbrace{(X\theta - y)^T}_{(m \times n)} \underbrace{(X\theta - y)}_{(n \times 1)} + \frac{1}{2} \sum_{i=1}^m \underbrace{(\theta^T x^{(i)} - y^{(i)})^2}_{(1 \times 1)}$$

$$\underbrace{\begin{bmatrix} - x^1 - \\ - x^2 - \\ \vdots \\ - x^m - \end{bmatrix}}_{(m \times n)} \underbrace{\begin{bmatrix} \theta_1 \\ \vdots \\ \theta_n \end{bmatrix}}_{(n \times 1)} = \begin{pmatrix} x^{(1)T} \theta \\ x^{(2)T} \theta \\ \vdots \\ x^{(m)T} \theta \end{pmatrix} - \begin{pmatrix} y^{(1)} \\ y^{(2)} \\ \vdots \\ y^{(m)} \end{pmatrix} =$$

Compute the gradient of $J(\theta)$:

$$\text{tr}(A) = \text{tr}(A^T)$$

$$\text{tr}(X) = X$$

$$\text{tr}(AB) = \text{tr}(BA)$$

$$\text{tr}(ABC) = \text{tr}(CAB) = \text{tr}(BCA)$$

$$\nabla_{\theta} J(\theta) = \nabla_{\theta} \left[\frac{1}{2} (X\theta - y)^T (X\theta - y) \right]$$

$$\nabla_A \text{tr}(AB) = B^T = \nabla_B \text{tr}(BA)$$

$$\nabla_x x^T A x = (A + A^T)x$$

$$\text{if } A = A^T, = 2Ax$$

$$A = X^T X$$

$$\begin{aligned} &= \frac{1}{2} \left(\nabla_{\theta} \left[\theta^T X^T X \theta - y^T X \theta - \theta^T X^T y + y^T y \right] \right) \\ &= \frac{1}{2} \left(\nabla_{\theta} \left[\theta^T X^T X \theta \right] - \nabla_{\theta} \left[y^T X \theta - \theta^T X^T y \right] + \nabla_{\theta} \left[y^T y \right] \right) \\ &= \frac{1}{2} \left(2X^T X \theta - \nabla_{\theta} \left(\text{tr}(y^T X \theta + \theta^T X^T y) \right) \right) \\ &\quad \text{tr}(y^T X \theta) + \text{tr}(\theta^T X^T y) \end{aligned}$$

$$= \frac{1}{2} (2X^T X \theta - 2X^T y)$$

$$= X^T X \theta - X^T y = 0$$

$$\theta = (X^T X)^{-1} X^T y$$

pseudo-inverse of X .

$$X\theta = y$$

$$\theta = X^+ y$$

ill-conditional

$$A^{-1}$$

conditional # of A \leftarrow max singular value

$$\kappa(A) = \frac{\delta_{\max}(A)}{\delta_{\min}(A)}$$

Compute the gradient of $J(\theta)$:

$$\begin{aligned}\nabla_{\theta} J(\theta) &= \nabla_{\theta} \left[\frac{1}{2} (\mathbf{X}\theta - \mathbf{y})^T (\mathbf{X}\theta - \mathbf{y}) \right] \\ &= \end{aligned}$$

Compute the gradient of $J(\theta)$:

$$\begin{aligned}\nabla_{\theta} J(\theta) &= \nabla_{\theta} \left[\frac{1}{2} (X\theta - y)^T (X\theta - y) \right] \\ &= X^T X\theta - X^T y\end{aligned}$$

Since $J(\theta)$ is **convex**, x is a global minimum of $J(\theta)$ when $\nabla J(\theta) = 0$.

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The Normal equation

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The Normal equation

$$\theta = (X^T X)^{-1} X^T y$$

$(X^T X)^{-1} X^T$ is called the **Moore-Penrose pseudoinverse of X**

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gradient descent	normal equation
iterative solution	exact solution

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Which method to use?

gradient descent	normal equation
iterative solution	exact solution
need to choose proper learning parameter α for cost function to converge	numerically unstable when X is ill-conditioned. e.g. features are highly correlated
works well for large number of samples m	solving equation is slow when m is large

Minimize $J(\theta)$ using Newton's Method

Newton's method solves real functions $f(x) = 0$ by iterative approximation

- ▶ Update rule: $x_{n+1} := x_n - \frac{f(x_n)}{f'(x_n)}$

Minimize $J(\theta)$ using Newton's Method

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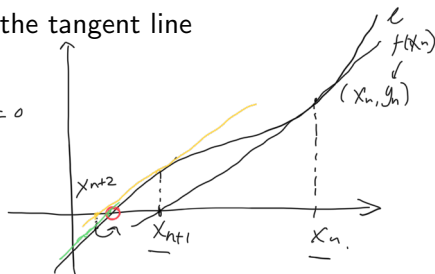
- ▶ Update rule: $x_{n+1} := x_n - \frac{f(x_n)}{f'(x_n)}$

Geometric intuition of Newton's method

- ▶ Find tangent line of f at (x_n, y_n)
- ▶ $x_{n+1} \leftarrow$ x-intercept of the tangent line
- ▶ $y_{n+1} \leftarrow f(x_{n+1})$

$$(\because y = f'(x_n)(x - x_n) + f(x_n) = 0$$

$$x = -\frac{f(x_n)}{f'(x_n)} + x_n$$



Newton's Method Demo

https://en.wikipedia.org/wiki/File:NewtonIteration_Ani.gif

Minimize $J(\theta)$ using Newton's Method

Newton's method for optimization $\min_{\theta} J(\theta)$

Use Newton's method to solve $\nabla_{\theta} J(\theta) = 0$:

- ▶ x is one-dimensional:

$$\theta := \theta - \frac{f'(x)}{f''(x)}$$

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$H(\theta) = \begin{bmatrix} \frac{\partial^2 J(\theta)}{\partial \theta^2} & \frac{\partial^2 J(\theta)}{\partial \theta_1 \partial \theta_2} & \dots & \frac{\partial^2 J(\theta)}{\partial \theta_1 \partial \theta_n} \\ \frac{\partial^2 J(\theta)}{\partial \theta_2 \partial \theta_1} & \dots & \dots & \dots \\ \vdots & \dots & \dots & \dots \\ \frac{\partial^2 J(\theta)}{\partial \theta_n \partial \theta_1} & \dots & \dots & \frac{\partial^2 J(\theta)}{\partial \theta_n^2} \end{bmatrix}$

► x is multidimensional:

$$\theta = \theta - \underbrace{H^{-1}(\theta)}_{\substack{\text{Hessian matrix} \\ \text{of } J(\theta)}} \underbrace{\nabla J(\theta)}_{\substack{\text{gradient} \\ \text{of } J(\theta)}}$$

where H is the Hessian matrix of $J(\theta)$.

a.k.a Newton-Raphson method

Newton's Method for Optimization

```
Initialize  $\theta$   
While  $\theta$  has not coveredged {  
   $\theta := \theta - H^{-1}(\theta)\nabla J(\theta)$   
}
```

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Performance of Newton's method:

- ▶ Needs fewer iterations than batch gradient descent

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Performance of Newton's method:

- ▶ Needs fewer iterations than batch gradient descent
- ▶ Computing H^{-1} is time consuming
- ▶ Faster in practice when n is small

Maximum Likelihood Estimation

Consider target y is modeled as

$$\underline{y}^{(i)} = \theta^T \underline{x}^{(i)} + \epsilon^{(i)}$$

and suppose $\epsilon^{(i)}$ are independently and identically distributed (IID)
to Gaussian distribution $\underline{\mathcal{N}}(\underline{0}, \underline{\sigma^2})$, IID.

Maximum Likelihood Estimation

Consider target y is modeled as

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$$p(\epsilon^{(i)}) =$$

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and suppose $\epsilon^{(i)}$ are *independently and identically distributed (IID)* to Gaussian distribution $\mathcal{N}(0, \sigma^2)$, then

$$p(\epsilon^{(i)}) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{\epsilon^{(i)2}}{2\sigma^2}\right)$$

Maximum Likelihood Estimation

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and suppose $\epsilon^{(i)}$ are *independently and identically distributed (IID)* to Gaussian distribution $\mathcal{N}(0, \sigma^2)$, then

$$p(\epsilon^{(i)}) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{\epsilon^{(i)2}}{2\sigma^2}\right)$$

deterministic parameter

$$p(\underbrace{y^{(i)}|x^{(i)}}_y; \theta) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y^{(i)} - \theta^T x^{(i)})^2}{2\sigma^2}\right)$$

Maximum Likelihood Estimation

The **likelihood** of this model with respect to θ is

$$L(\theta) = \underbrace{p(\vec{y}|X; \theta)} = \prod_{i=1}^m \underbrace{p(y^{(i)}|x^{(i)}; \theta)}$$

Maximum Likelihood Estimation

The **likelihood** of this model with respect to θ is

$$L(\theta) = p(\vec{y}|X; \theta) = \prod_{i=1}^m p(y^{(i)}|x^{(i)}; \theta)$$

Maximum likelihood estimation of θ :

$$\theta_{MLE} = \operatorname{argmax}_{\theta} L(\theta)$$

Maximum Likelihood Estimation

We compute log likelihood,

$$\begin{aligned} \log L(\theta) &= \log \prod_{i=1}^m p(y^{(i)} | x^{(i)}; \theta) = \sum_{i=1}^m \log p(y^{(i)} | x^{(i)}; \theta) \\ &= \sum_{i=1}^m \log \left(\frac{1}{\sqrt{2\pi\sigma^2}} \exp \left(-\frac{(y^{(i)} - \theta^T x^i)^2}{2\sigma^2} \right) \right) \\ &= \sum_{i=1}^m \log \frac{1}{\sqrt{2\pi\sigma^2}} + \left(-\frac{(y^i - \theta^T x^i)^2}{2\sigma^2} \right) \\ &= \left(\sum_{i=1}^m \log \frac{1}{\sqrt{2\pi\sigma^2}} - \frac{1}{\sigma^2} \frac{1}{2} (y^i - \theta^T x^i)^2 \right) \\ &= m \log \frac{1}{\sqrt{2\pi\sigma^2}} - \frac{1}{2\sigma^2} \sum_{i=1}^m (y^i - \theta^T x^i)^2 \end{aligned}$$

$$\max_{\theta} L(\theta) \leftrightarrow \min_{\theta} J(\theta) \quad \leftarrow \text{least square}$$

Maximum Likelihood Estimation

We compute log likelihood,

$$\begin{aligned}\log L(\theta) &= \log \prod_{i=1}^m p(y^{(i)}|x^{(i)}; \theta) = \sum_{i=1}^m \log p(y^{(i)}|x^{(i)}; \theta) \\ &= \sum_{i=1}^m \log \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y^{(i)} - \theta^T x)^2}{2\sigma^2}\right)\end{aligned}$$

Maximum Likelihood Estimation

We compute log likelihood,

$$\begin{aligned}\log L(\theta) &= \log \prod_{i=1}^m p(y^{(i)}|x^{(i)}; \theta) = \sum_{i=1}^m \log p(y^{(i)}|x^{(i)}; \theta) \\ &= \sum_{i=1}^m \log \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y^{(i)} - \theta^T x)^2}{2\sigma^2}\right) \\ &= m \log \frac{1}{\sqrt{2\pi\sigma^2}} - \frac{1}{\sigma^2} \cdot \frac{1}{2} \sum_{i=1}^m (y^{(i)} - \theta^T x^{(i)})^2\end{aligned}$$

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Then $\operatorname{argmax}_{\theta} \log L(\theta) \equiv \operatorname{argmin}_{\theta} \frac{1}{2} \sum_{i=1}^m (y^{(i)} - \theta^T x^{(i)})^2$.

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Under the assumptions on $\epsilon^{(i)}$, least-squares regression corresponds to the maximum likelihood estimate of θ .

Linear Regression Summary

- ▶ Least square regression
- ▶ Solving least square:
 - ▶ gradient descent
 - ▶ normal equation
 - ▶ newton's method
- ▶ Probabilistic interpretation: maximum likelihood

Logistic Regression

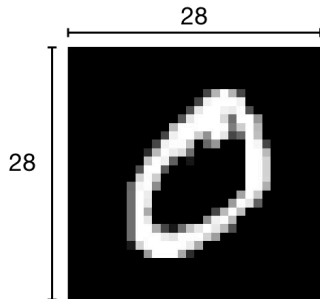
A binary classification problem

Classify binary digits

- ▶ Training data: 12600 grayscale images of handwritten digits



- ▶ Each image is represented by a vector $x^{(i)}$ of dimension $28 \times 28 = 784$
- ▶ Vectors $x^{(i)}$ are normalized to $[0,1]$



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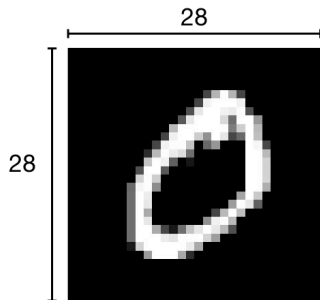
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Binary classification: $\mathcal{Y} = \{0, 1\}$

- ▶ negative class: $y^{(i)} = 0$
- ▶ positive class: $y^{(i)} = 1$

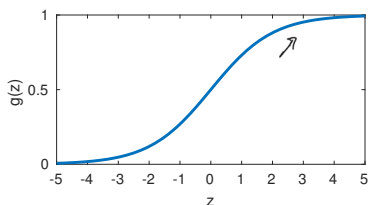


Logistic Regression Hypothesis Function

Sigmoid function

$$g(z) = \frac{1}{1 + e^{-z}}$$

- ▶ $g : \mathbb{R} \rightarrow (0, 1)$
- ▶ $g'(z) = g(z)(1 - g(z))$

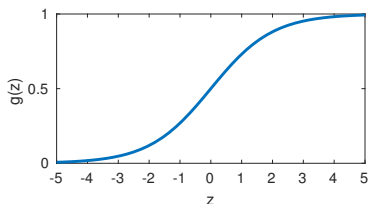


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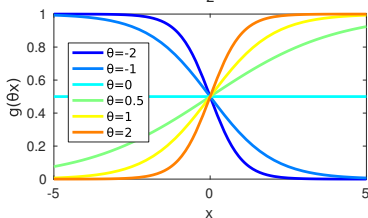
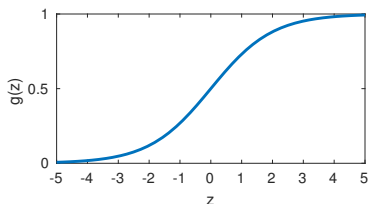


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Hypothesis function for logistic regression: z *logistic.*

$$h_{\theta} = g(\underbrace{\theta^T x}_z) = \frac{1}{\underbrace{1 + e^{-\theta^T x}}}$$

Maximum likelihood estimation for logistic regression

Logistic regression assumes $y|x$ is **Bernoulli distributed**.

e.g. tossing a coin with $p(\text{head}) = h_{\theta}(x)$

- ▶ $p(y = 1 | x; \theta) = \boxed{h_{\theta}(x)} = \phi$ $y|x \sim \text{Bernoulli}(\phi)$.
- ▶ $p(y = 0 | x; \theta) = 1 - \underbrace{h_{\theta}(x)} = 1 - \phi$.
 $p(\text{tail}) \nearrow$

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▶ $p(y = 0 | x; \theta) = 1 - h_{\theta}(x)$

$$p(y | x; \theta) = \underbrace{(h_{\theta}(x))^y}_{y=0 \uparrow 0} \underbrace{(1 - h_{\theta}(x))^{1-y}}_{1 - h_{\theta}(x)}$$

(Note: In the original image, there are handwritten annotations: a circled 'y' with an arrow pointing to '1', and a circled '1-y' with an arrow pointing to '1'.)

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▶ $p(y = 0 | x; \theta) = 1 - h_{\theta}(x)$

$$p(y | x; \theta) = (h_{\theta}(x))^y (1 - h_{\theta}(x))^{1-y}$$

Given m independently generated training examples, the likelihood function is:

$$L(\theta) = p(\bar{y}|X; \theta) = \prod_{i=1}^m p(y^{(i)}|x^{(i)}; \theta) = \log \prod_{i=1}^m h_{\theta}(x^{(i)})^{y^{(i)}} (1 - h_{\theta}(x^{(i)}))^{1-y^{(i)}}.$$

$$l(\theta) = \log(L(\theta)) = \sum_{i=1}^m \underbrace{y^{(i)} \log h_{\theta}(x^{(i)})}_{\text{}} + \underbrace{(1 - y^{(i)}) \log(1 - h_{\theta}(x^{(i)}))}_{\text{}}.$$

Maximum likelihood estimation for logistic regression

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e.g. tossing a coin with $p(\text{head}) = h_{\theta}(x)$

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Given m independently generated training examples, the likelihood function is:

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$$l(\theta) = \log(L(\theta)) = \sum_{i=1}^m y^{(i)} \log h_{\theta}(x^{(i)}) + (1 - y^{(i)}) \log(1 - h_{\theta}(x^{(i)}))$$

$l(\theta)$ is concave!

Maximum likelihood estimation for logistic regression

$$l(\theta) = \sum_{i=1}^m y^{(i)} \log h_{\theta}(x^{(i)}) + (1 - y^{(i)}) \log(1 - h_{\theta}(x^{(i)}))$$

Solve $\operatorname{argmax}_{\theta} l(\theta)$ using gradient ascent:

$$\begin{aligned}
 g'(x) &= g(x)(1-g(x)) & \frac{\partial l(\theta)}{\partial \theta_j} &= \sum_{i=1}^m y^i \frac{1}{h_{\theta}(x^i)} \left[\frac{\partial}{\partial \theta_j} h_{\theta}(x^i) \right] + \frac{1-y^i}{1-h_{\theta}(x^i)} \left[\frac{\partial}{\partial \theta_j} (-1) h_{\theta}(x^i) \right] \\
 h_{\theta}(x^i) &= g(\theta^T x^i) & &= \sum_{i=1}^m \left[y^i \frac{1}{h_{\theta}(x^i)} - (1-y^i) \frac{1}{1-h_{\theta}(x^i)} \right] \frac{\partial}{\partial \theta_j} h_{\theta}(x^i) \\
 \frac{\partial}{\partial \theta_j} g(\theta^T x^i) &= \underbrace{g(\theta^T x^i)} \underbrace{(1-g(\theta^T x^i))}_{x_j^i} \frac{\partial}{\partial \theta_j} \theta^T x^i & &= \sum_{i=1}^m \left(\frac{y^i h_{\theta}(x^i) (1-h_{\theta}(x^i)) x_j^i}{h_{\theta}(x^i)} - \frac{(1-y^i) h_{\theta}(x^i) (1-h_{\theta}(x^i)) x_j^i}{1-h_{\theta}(x^i)} \right) \\
 &= \underbrace{h_{\theta}(x^i) (1-h_{\theta}(x^i))}_{x_j^i} x_j^i & &= \sum_{i=1}^m (y^i (1-h_{\theta}(x^i)) - (1-y^i) h_{\theta}(x^i)) x_j^i \\
 & & &= \sum_{i=1}^m (y^i - h_{\theta}(x^i)) x_j^i \\
 & & & \swarrow \quad \searrow \\
 & & & \theta^T x \quad \underline{g(\theta^T x)}
 \end{aligned}$$

Maximum likelihood estimation for logistic regression

$$l(\theta) = \sum_{i=1}^m y^{(i)} \log h_{\theta}(x^{(i)}) + (1 - y^{(i)}) \log(1 - h_{\theta}(x^{(i)}))$$

Solve $\operatorname{argmax}_{\theta} l(\theta)$ using gradient ascent:

$$\frac{\partial l(\theta)}{\partial \theta_j} = \sum_{i=1}^m (y^{(i)} - h_{\theta}(x^{(i)})) x_j^{(i)}$$

Stochastic Gradient Ascent

```
Repeat until convergence {
  for  $i = 1 \dots m$  {
     $\theta_j = \theta_j + \alpha (y^{(i)} - h_{\theta}(x^{(i)})) x_j^{(i)}$  for every  $j$ 
  }
}
```

- Update rule has the same form as least square regression, but with different hypothesis function h_{θ}

Binary Digit Classification

Using the learned classifier

Given an image x , the predicted label is

$$\hat{y} = \begin{cases} 1 & g(\theta^T x) > 0.5 \\ 0 & \text{otherwise} \end{cases}$$

Binary digit classification results

	sample size	accuracy
Training	16200	100%
Testing	1225	100%

- ▶ Testing accuracy is 100% since this problem is relatively easy.

Multi-Class Classification

Multiple Binary Classifiers

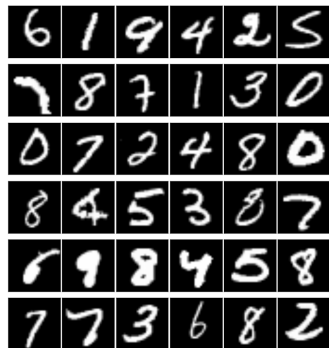
Softmax Regression

Multi-class classification

Each data sample belong to one of $k > 2$ different classes.

$$\mathcal{Y} = \{1, \dots, k\}$$

MNIST Samples $|\mathcal{Y}| = 10.$



Given new sample $x \in \mathbb{R}^k$, predict which class it belongs.

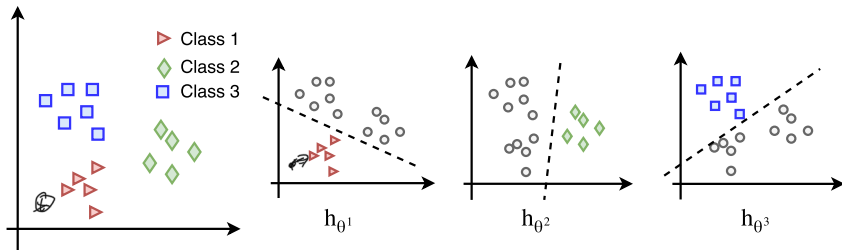
Naive Approach: Convert to binary classification

One-Vs-Rest

Learn k classifiers h_1, \dots, h_k . Each h_i classify one class against the rest of the classes.

Given a new data sample x , its predicted label \hat{y} :

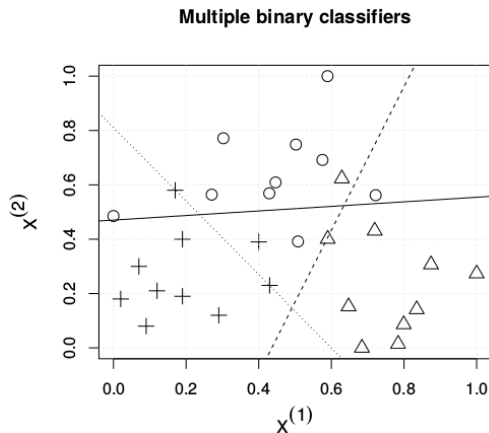
$$\hat{y} = \underset{\underline{1..k}}{\operatorname{argmax}}_i \underbrace{h_i(x)}$$



Multiple binary classifiers

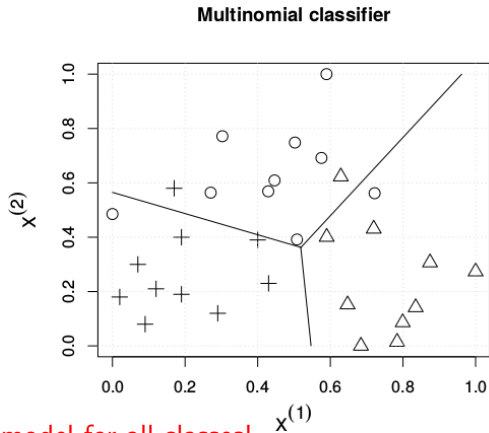
Drawbacks of One-Vs-Rest:

- ▶ Class unbalance: more negative samples than positive samples
- ▶ Different classifiers may have different confidence scales



Drawbacks of One-Vs-Rest:

- ▶ Class imbalance: more negative samples than positive samples
- ▶ Different classifiers may have different confidence scales



Learn one model for all classes!

Extend logistic regression: Softmax Regression

Assume $p(y|x)$ is **multinomial distributed**.

e.g. outcomes of rolling a k -sided die n times, each side has independent probability ϕ_1, \dots, ϕ_k

$$\underbrace{\phi_1, \dots, \phi_k}_{h_{\theta}(x)}$$

Extend logistic regression: Softmax Regression

Assume $p(y|x)$ is **multinomial distributed**.

e.g. outcomes of rolling a k -sided die n times, each side has independent probability ϕ_1, \dots, ϕ_k

Hypothesis function for sample x :

$$h_{\theta}(x) = \begin{matrix} k \\ \left[\begin{array}{c} p(y=1|x;\theta) \\ \vdots \\ p(y=k|x;\theta) \end{array} \right] \end{matrix} = \frac{1}{\sum_{j=1}^k e^{\theta_j^T x}} \begin{bmatrix} e^{\theta_1^T x} \\ \vdots \\ e^{\theta_k^T x} \end{bmatrix} = \text{softmax}(\theta^T x)$$

$g(z)$

$$\text{softmax}(z_j) = \frac{e^{z_j}}{\sum_{j=1}^k e^{z_j}}$$

Extend logistic regression: Softmax Regression

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$$\text{softmax}(z_i) = \frac{e^{z_i}}{\sum_{j=1}^k e^{z_j}}$$

Parameters: $\theta = \begin{bmatrix} - & \theta_1^T & - \\ \vdots & & \\ - & \theta_k^T & - \end{bmatrix}$ $\left. \vphantom{\begin{bmatrix} - & \theta_1^T & - \\ \vdots & & \\ - & \theta_k^T & - \end{bmatrix}} \right\} k \times n$

Softmax Regression

Given $(x^{(i)}, y^{(i)})$, $i = 1, \dots, m$, the log-likelihood of the Softmax model is

$$\begin{aligned} \ell(\theta) &= \sum_{i=1}^m \log p(y^{(i)} | x^{(i)}; \theta) \\ &= \sum_{i=1}^m \log \prod_{l=1}^k p(y^{(i)} = l | x^{(i)}) \mathbf{1}_{\{y^{(i)}=l\}} \end{aligned} \quad \begin{array}{l} \downarrow \\ \text{1} \quad y^i = l. \\ \text{0} \quad y^i \neq l. \end{array}$$

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Softmax Regression

Derive the stochastic gradient descent update:

- ▶ Find $\nabla_{\theta_l} \ell(\theta)$

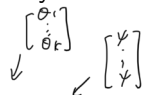
$$\nabla_{\theta_l} \ell(\theta) = \sum_{i=1}^m \left[\left(\mathbf{1}\{y^{(i)} = l\} - P(y^{(i)} = l | x^{(i)}; \theta) \right) x^{(i)} \right]$$

Property of Softmax Regression

$$P(y^i = l | x^i) = \frac{e^{\theta_l^T x^i}}{\sum_{j=1}^k e^{\theta_j^T x^i}}$$

- ▶ Parameters $\theta_1, \dots, \theta_k$ are not independent:
 $\sum_j p(y = j | x) = \sum_j \phi_j = 1$
- ▶ Knowing $k - 1$ parameters completely determines model.

Invariant to scalar addition



Proof.
$$p(y=l|x; \theta - \psi) = \frac{e^{(\theta - \psi)^T x}}{\sum_{j=1}^k e^{(\theta_j - \psi_j)^T x}} = \frac{e^{\theta^T x} \cdot (e^{-\psi^T x})}{\sum_{j=1}^k e^{\theta_j^T x} \cdot (e^{-\psi_j^T x})} = \frac{e^{-\psi^T x}}{e^{-\psi^T x}} \cdot \frac{e^{\theta^T x}}{\sum_{j=1}^k e^{\theta_j^T x}} = P(y^i = l | x^i)$$

Relationship with Logistic Regression

When $K = 2$,

$$h_{\theta}(x) = \frac{1}{e^{\theta_1^T x} + e^{\theta_2^T x}} \begin{bmatrix} e^{\theta_1^T x} \\ e^{\theta_2^T x} \end{bmatrix}$$

Relationship with Logistic Regression

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$$h_{\theta}(x) = \frac{1}{e^{\theta_1^T x} + e^{\theta_2^T x}} \begin{bmatrix} e^{\theta_1^T x} \\ e^{\theta_2^T x} \end{bmatrix}$$

Replace $\theta = \begin{bmatrix} \theta_1 \\ \theta_2 \end{bmatrix}$ with $\theta - \begin{bmatrix} \theta_2 \\ \theta_2 \end{bmatrix} = \begin{bmatrix} \theta_1 - \theta_2 \\ 0 \end{bmatrix} = \tilde{\theta}$ since $p(y|x;\theta)$ is invariant to scalar additions

$$h_{\theta}(x) = \frac{1}{e^{\theta_1^T - \theta_2^T} x + e^{0^T x}} \begin{bmatrix} e^{(\theta_1 - \theta_2)^T x} \\ e^{0^T x} \end{bmatrix}$$

$$= \begin{bmatrix} \frac{e^{(\theta_1 - \theta_2)^T x}}{1 + e^{(\theta_1 - \theta_2)^T x}} \\ \frac{1}{1 + e^{(\theta_1 - \theta_2)^T x}} \end{bmatrix} = \begin{bmatrix} \frac{1}{1 + e^{-(\theta_1 - \theta_2)^T x}} \\ 1 - \frac{1}{1 + e^{-(\theta_1 - \theta_2)^T x}} \end{bmatrix} = \begin{bmatrix} g(\tilde{\theta}^T x) \\ 1 - g(\tilde{\theta}^T x) \end{bmatrix}$$

When to use Softmax?

OH : 6:30 - 8:30

Rm 1108A.

- ▶ When classes are mutually exclusive: use Softmax
- ▶ Not mutually exclusive: multiple binary classifiers may be better

multi-label classification

↑
when labels are not mutually exclusive.